Perturbation Methods for Markov-Switching Models*

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Abstract

Markov Switching models are a way to consider discrete changes in the economic environment, such as policy changes, and allow agents in the economy to form expectations over these changes. This paper develops a methodology for constructing approximations to the solution of Markov Switching dynamic stochastic general equilibrium (MS-DSGE) models. The method allows for changes in parameters that both do and do not affect the economy’s steady state, and enables linear or higher-order approximations. In addition, the paper proves that first-order approximations to a wide class of MS-DSGE models are not certainty equivalent. The numerical procedure handles potentially large systems and considers existence and uniqueness using the concept of mean square stability. Two examples, one Real Business Cycle and one New Keynesian, illustrate the procedure and issues of certainty equivalence and mean square stability.

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1 Introduction

Following the introduction of vector autoregressions (VARs) to macroeconomics by it was quickly realized that it is difficult to find macroeconomic applications for which model parameters remain stable over long periods of time. This problem was not unique to reduced form representations of the data, but was also an issue when more structural approaches were considered. One way to solve the problem, pursued by and followed up by, breaks the sample into sub-periods and estimates the structural models in which one or more of the model’s parameters differ across sub-samples. While this approach addresses the parameter instability problem, it fails to consider that forward looking agents living in a world in which parameters are known to change occasionally would be expected to take possible parameter change into account when forming their expectations and, therefore, will affect their optimal decisions.

An alternative approach to parameter instability, suggested by the work of and pursued in, is to estimate a backward-looking vector autoregression (VAR) with regime dependent parameters. This approach has its limitations since it does not allow for the presence of forward-looking components that are present in a dynamic stochastic general equilibrium (DSGE) model.

A number of authors have recently studied forward looking Markov-switching linear rational expectations (MSLRE) models. Work in this area includes papers by, and. MSLRE models are more complicated than linear rational expectations models since the agents of the model must be allowed to take account of the possibility of future regime changes when forming expectations. The MSLRE literature has made some headway in addressing questions like setting necessary and sufficient conditions to determine if the parameters of a Markov-switching rational expectations model lead to a determinate equilibrium (See).

There are two main shortcomings with the MSLRE approach. First, most of the analyzed models do not begin from first principles. In other words, researchers consider linear rational expectations (LRE) models where Markov-switching (MS) has been added after the model has been linearized. Second, higher order solutions are not considered. Given that MS parameters add a lot of uncertainty to the model, considering higher order approximations may be potentially important. This paper solves these two shortcomings. In particular, it shows how to use
perturbation methods to solve Markov-switching rational expectations (MSRE) models - note the absence of the “linear” - starting from first principles, i.e. from the set of (non-linearized) first order conditions that define equilibrium.

Following ?, ?, and ?, this paper uses the concept of mean square stability (MSS) to characterize stable solutions. The perturbation approach uses the theory of Gröbner Bases to find solutions, and determines existence and uniqueness of MSS solutions. It also allows for a flexible regime-switching specification, including in parameters that affect the steady state of the economy. In particular, the first order approximation of models where switching affects the steady state is not certainty equivalent.

After developing the methodology, the paper presents two example economies that illustrate the methodology and highlight the issues of mean square stability and certainty equivalence. In the first, a simple real business cycle model with stochastic drift shows how to use the methodology and the importance of certainty equivalence. The second, a New Keynesian model, adds sticky prices and a monetary authority with changes in the policy rule, and shows how mean square stability determines existence and uniqueness.

The remainder of the paper is as follows: Section 2 describes a general class of MS-DSGE models and the nature of Markov switching. Sections 3 and 4 discuss the first-order approximation, the former showing how to solve the model, and the latter highlighting the key quadratic equations and how to use Gröbner Bases to solve them. Section 6.1 has an example RBC economy, Section 6.6 has an example NK economy, and Section 7 concludes.

2 The Model

Consider a dynamic general equilibrium model in which some of the parameters follow a discrete state Markov chain indexed by $s_t$ with transition matrix $P = (p_{s,s'})$. The element $p_{s,s'}$ represents the probability that $s_{t+1} = s'$ given $s_t = s$ for $s, s' \in \{1, \ldots, n_s\}$ where $n_s$ is the number of regimes and when $s_t = s$ the model is said to be in regime $s$ at time $t$. The vector of changing
parameters $\theta_t$ has size $n_{\theta} \times 1$. Given any $x_{t-1}, \varepsilon_t$, and $\theta_t$, the set of equilibrium conditions of a wide class of models can be written as

$$\mathbb{E}_t f (y_{t+1}, y_t, x_t, x_{t-1}, \chi \varepsilon_{t+1}, \varepsilon_t, \theta_{t+1}, \theta_t) = 0_{n_x+n_y}$$

(1)

where $\mathbb{E}_t$ denotes the mathematical expectations operator conditional on information available at time $t$, and $0_{m_1 \times m_2}$ denotes a $m_1 \times m_2$ matrix of zeros. The vector $x_{t-1}$ of predetermined variables (endogenous and exogenous) is of size $n_x \times 1$, the vector $y_t$ of non-predicted variables is of size $n_y \times 1$, the vector $\varepsilon_t$ of independent innovations to the exogenous predetermined variables with mean equal to zero is of size $n_\varepsilon \times 1$, and $\chi$ is the perturbation parameter. The function $f$ maps $\mathbb{R}^{2(n_y+n_x+n_\varepsilon+n_\theta)}$ into $\mathbb{R}^{n_y+n_x}$ is the number of equations in (1). Since the parameters, $\theta_t$, in (1) depend on the state of the Markov chain, there are $n_s$ sets of equilibrium conditions, one for each value of the Markov chain, instead of the single set of equilibrium conditions in the constant parameter case.

The solution to the model has the form

$$y_t \equiv g (x_{t-1}, \varepsilon_t, \chi, s_t),$$

(2)

$$y_{t+1} \equiv g (x_t, \chi \varepsilon_{t+1}, \chi, s_{t+1}),$$

(3)

and

$$x_t \equiv h (x_{t-1}, \varepsilon_t, \chi, s_t)$$

(4)

where $g$ maps $\mathbb{R}^{n_x+n_\varepsilon+1} \times \{1, \ldots, n_s\}$ into $\mathbb{R}^{n_y}$ and $h$ maps $\mathbb{R}^{n_x+n_\varepsilon+1} \times \{1, \ldots, n_s\}$ into $\mathbb{R}^{n_x}$. The goal is to find the Taylor expansion of the functions $g$ and $h$ around the steady state.

The parameters $\theta_t$ depend on the regime in the following way

$$\theta_t \equiv \theta (\chi, s_t) \quad \text{and} \quad \theta_{t+1} \equiv \theta (\chi, s_{t+1})$$

(5)

where $\theta$ maps $\mathbb{R} \times \{1, \ldots, n_s\}$ into $\mathbb{R}^{n_\theta}$. The vector of parameters $\theta_t$ has two subvectors $\theta_{1t}$ and $\theta_{2t}$

$$\theta_t = \left( \theta'_{1t} \quad \theta'_{2t} \right)' \equiv \left( \theta_1 (\chi, s_t)' \quad \theta_2 (\chi, s_t)' \right)'$$

(6)

$\footnotetext{There may also be a set of non-changing parameters not included in $\theta_t$.}$
where $\theta_{1t}$ and $\theta_{2t}$ have sizes $n_{\theta_1}$ and $n_{\theta_2}$, respectively, and

$$\theta_1(\chi, s_t) = \bar{\theta}_1 + \chi \theta_1(s_t) \quad (7)$$

and

$$\theta_2(\chi, s_t) = \hat{\theta}_2(s_t). \quad (8)$$

The parameters $\theta_{t+1}$ have the same functional forms.\(^2\) Note two things about this specification: first, $\hat{\theta}_1(s_t)$ is the deviation of $\theta_{1t}$ from $\bar{\theta}_1$ in regime $s_t$ and, second, $\theta_{2t}$ is not a function of the perturbation parameter $\chi$. Hence, the perturbation parameter, $\chi$, only affects a subset of the parameters, $\theta_{1t}$, while $\theta_{2t}$ is not affected by the perturbation parameter. The choice of which parameters to perturb, $\theta_{1t}$, and which ones do not perturb, $\theta_{2t}$, is not unique, but there is one restriction. Define the steady state of the model as vectors $x_{ss}$ and $y_{ss}$ of size $n_x \times 1$ and $n_y \times 1$ respectively such that

$$f \left( y_{ss}, x_{ss}, x_{ss}, 0_{n_x}, 0_{n_x}, \left( \bar{\theta}_1', \hat{\theta}_2(s_{t+1})' \right)', \left( \bar{\theta}_1', \hat{\theta}_2(s_t)' \right)' \right) = 0_{n_x+n_y}$$

for all $s_{t+1}$ and $s_t$. Thus, the partition should be such that neither $\theta_2(0, s_{t+1}) = \hat{\theta}_2(s_{t+1})$ nor $\theta_2(0, s_t) = \hat{\theta}_2(s_t)$ enter in the calculation of the steady state since the last expression has to hold for all $s_{t+1}$ and $s_t$.

In other words, the partition should be such that the function $f$, once $\theta_{1t}$ and $\theta_{2t}$ have been replaced by (7) and (8) and evaluated at $\varepsilon_t = 0_{n_x}$ and $\chi = 0$, can be written as another function $f_{ss}$ that only depends on $y_{t+1}, y_t, x_t, x_{t-1}$, and $\bar{\theta}_1$ but neither on $\hat{\theta}_2(s_{t+1})$ nor $\hat{\theta}_2(s_t)$, i.e.

$$f \left( y_{t+1}, y_t, x_t, x_{t-1}, 0_{n_x}, 0_{n_x}, \left( \bar{\theta}_1', \hat{\theta}_2(s_{t+1})' \right)', \left( \bar{\theta}_1', \hat{\theta}_2(s_t)' \right)' \right) = f_{ss} \left( y_{t+1}, y_t, x_t, x_{t-1}, \bar{\theta}_1, \bar{\theta}_1 \right)$$

where $0_m$ is a vector of zeros of size $n_m \times 1$.

In general, more than one partition of parameters accomplishes this objective. In any case, included in $\theta_{1t}$ is the smallest set of parameters such that the steady state is defined as described above. Since the steady state depends upon $\bar{\theta}_1$, a natural choice for this point is the mean of

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\(^2\)These functional forms are not necessary but just convenient for the derivations; any other functional form such that $\theta_1(0, s_t) = \bar{\theta}_1$ for all $s_t$ holds may also work.
the ergodic distribution across $\theta_{1t}$; but again, this selection is not unique. Sections 6 provides examples of partitions of $\theta_t$ and choices of $\bar{\theta}_1$.

Given the definition of the steady state, it is the case that

$$y_{ss} = g(x_{ss}, 0_{n_s}, 0, s_t) \text{ and } x_{ss} = h(x_{ss}, 0_{n_s}, 0, s_t)$$

for all $s_t$ and

$$y_{ss} = g(x_{ss}, 0_{n_s}, 0, s_{t+1}) \text{ and } x_{ss} = h(x_{ss}, 0_{n_s}, 0, s_{t+1})$$

for all $s_{t+1}$.

Using equations (2), (3), (4), and (5) the function $f$ can be written as

$$F(x_{t-1}, \varepsilon_{t+1}, \varepsilon_t, s_{t+1}, \chi, s_t) = f\left( g(x_{t-1}, \varepsilon_t, \chi, s_t), \varepsilon_{t+1}, \chi, s_{t+1} \right)$$

for all $x_{t-1}, \varepsilon_{t+1}, \varepsilon_t, s_{t+1}$, and $s_t$. The function $F$ maps $\mathbb{R}^{n_x+n_z+1} \times \{1, \ldots, n_s\} \times \{1, \ldots, n_s\}$ into $\mathbb{R}^{n_y+n_x}$.

Assuming that innovations to the exogenous predetermined variables, $\varepsilon_t$, are independent of the Markov chain, $s_t$, write (1) as

$$\mathbb{E}_t f(y_{t+1}, y_t, x_t, x_{t-1}, \varepsilon_{t+1}, \varepsilon_t, \theta_{t+1}, \theta_t) = \mathbb{G}(x_{t-1}, \varepsilon_t, \chi, s_t) = \sum_{s'=1}^{n_s} p_{st,s'} \int F(x_{t-1}, \varepsilon', \varepsilon_t, s', \chi, s_t) \mu(\varepsilon') d\varepsilon' = 0_{n_y+n_x}$$

for all $x_{t-1}, \varepsilon_t$, and $s_t$ where $\mu$ is the density of the innovations. The function $\mathbb{G}$ maps $\mathbb{R}^{n_x+n_z+1} \times \{1, \ldots, n_s\}$ into $\mathbb{R}^{n_y+n_x}$.

The remainder of the paper will use the following notation

$$\mathcal{D}G_t(x_{t-1}, \varepsilon_t, \chi, s_t) = [\mathcal{D}G_t^i(x_{t-1}, \varepsilon_t, \chi, s_t)]_{1 \leq i \leq n_y+n_z, 1 \leq j \leq n_x+n_z+1}$$

to refer the $(n_y+n_x) \times (n_x+n_z+1)$ matrix of partial derivatives of $G$ with respect to $(x_{t-1}, \varepsilon_t, \chi)$ evaluated at $(x_{t-1}, \varepsilon_t, \chi, s_t)$ for all $x_{t-1}, \varepsilon_t$, and $s_t$. Note the absence of derivatives with respect to $s_t$, since it is a discrete variable. Equivalently,

$$\mathcal{D}G(x_{ss}, 0_{n_s}, 0, s_t) = [\mathcal{D}G^i(x_{ss}, 0_{n_s}, 0, s_t)]_{1 \leq i \leq n_y+n_z, 1 \leq j \leq n_x+n_z+1}$$
refers to the \((n_y + n_x) \times (n_x + n_\varepsilon + 1)\) matrix of partial derivatives of \(G\) with respect to \((x_{t-1}, \varepsilon_t, \chi)\) evaluated at \((x_{ss}, 0_n, 0, s_t)\) for all \(s_t\). To simplify notation define

\[
\mathcal{D}G_{ss} (s_t) \equiv \mathcal{D}G (x_{ss}, 0_n, 0, s_t) \quad \text{and} \quad \mathcal{D}jG_{ss}^{i} (s_t) \equiv \mathcal{D}jG^{i} (x_{ss}, 0_n, 0, s_t)
\]

for all \(i, j, \) and \(s_t\). Thus,

\[
\mathcal{D}G_{ss} (s_t) = [\mathcal{D}jG_{ss}^{i} (s_t)]_{1 \leq i \leq n_y+n_x, 1 \leq j \leq n_x+n_\varepsilon+1}
\]

for all \(s_t\). In the same way,

\[
\mathcal{D}f_{ss} (s_{t+1}, s_t) = \\
\left[ \mathcal{D}jJ^{i} \left( y_{ss}, y_{ss}, x_{ss}, x_{ss}, 0_n, 0_n; \left( \hat{\theta}_1, \hat{\theta}_2 \right)(s_{t+1})' \right) \right]_{1 \leq i \leq n_y+n_x, 1 \leq j \leq 2(n_y+n_x+n_\varepsilon+n_\theta)}^{1 \leq i \leq n_y+n_x, 1 \leq j \leq 2(n_y+n_x+n_\varepsilon+n_\theta)}
\]

is the \((n_y + n_x) \times (2(n_y + n_x + n_\varepsilon + n_\theta))\) matrix of partial derivatives of \(f\) with respect to all its components evaluated at \(\left( y_{ss}, y_{ss}, x_{ss}, x_{ss}, 0_n, 0_n; \left( \hat{\theta}_1, \hat{\theta}_2 \right)(s_{t+1})' \right)\) for all \(s_{t+1}\) and \(s_t\),

\[
\mathcal{D}g (x_{ss}, 0_n, 0, s_t) = [\mathcal{D}jG^{i} (x_{ss}, 0_n, 0, s_t)]_{1 \leq i \leq n_y, 1 \leq j \leq n_x+n_\varepsilon+1}
\]

is the \(n_y \times (n_x + n_\varepsilon + 1)\) matrix of partial derivatives of \(g\) with respect to \((x_{t-1}, \varepsilon_t, \chi)\) evaluated at \((x_{ss}, 0_n, 0, s_t)\) for all \(s_t\), and

\[
\mathcal{D}h (x_{ss}, 0_n, 0, s_t) = [\mathcal{D}jH^{i} (x_{ss}, 0_n, 0, s_t)]_{1 \leq i \leq n_x, 1 \leq j \leq n_x+n_\varepsilon+1}
\]

is the \(n_x \times (n_x + n_\varepsilon + 1)\) matrix of partial derivatives of \(h\) with respect to \((x_{t-1}, \varepsilon_t, \chi)\) evaluated at \((x_{ss}, 0_n, 0, s_t)\) for all \(s_t\). To simplify notation, define

\[
\mathcal{D}g_{ss}^{i} (s_t) \equiv \mathcal{D}g^{i} (x_{ss}, 0_n, 0, s_t) \quad \text{and} \quad \mathcal{D}jg_{ss}^{i} (s_t) \equiv \mathcal{D}jG^{i} (x_{ss}, 0_n, 0, s_t)
\]

for all \(i, j, \) and \(s_t\) and

\[
\mathcal{D}h_{ss}^{i} (s_t) \equiv \mathcal{D}h^{i} (x_{ss}, 0_n, 0, s_t) \quad \text{and} \quad \mathcal{D}jh_{ss}^{i} (s_t) \equiv \mathcal{D}jH^{i} (x_{ss}, 0_n, 0, s_t)
\]

for all \(i, j, \) and \(s_t\). Thus,

\[
\mathcal{D}g_{ss} (s_t) = [\mathcal{D}jg_{ss}^{i} (s_t)]_{1 \leq i \leq n_y+n_x, 1 \leq j \leq n_x+n_\varepsilon+1}
\]
and
\[ D h_{ss}(s_t) = [D_j h_{ss}^i(s_t)]_{1 \leq i \leq n_y + n_x, 1 \leq j \leq n_x + n_s + 1} \]
for all \( s_t \).

### 3 First Order Approximation

This section shows how to find the first order Taylor expansions to \( g \) and \( h \) around the point \((x_{ss}, 0_{n_x}, 0, s_t)\) of the form
\[
g(x_{t-1}, \varepsilon_t, \chi, s_t) - y_{ss} \simeq [D_1 g_{ss}(s_t), \ldots, D_{n_x} g_{ss}(s_t)] (x_{t-1} - x_{ss}) \\
+ [D_{n_x+1} g_{ss}(s_t), \ldots, D_{n_x+n_s} g_{ss}(s_t)] \varepsilon_t + D_{n_x+n_s+1} g_{ss}(s_t) \chi
\]
and
\[
h(x_{t-1}, \varepsilon_t, \chi, s_t) - x_{ss} \simeq [D_1 h_{ss}(s_t), \ldots, D_{n_x} h_{ss}(s_t)] (x_{t-1} - x_{ss}) \\
+ [D_{n_x+1} h_{ss}(s_t), \ldots, D_{n_x+n_s} h_{ss}(s_t)] \varepsilon_t + D_{n_x+n_s+1} h_{ss}(s_t) \chi
\]
for all \( s_t \) where \( D_j g_{ss}(s_t) \) is the \( j^{th} \) column vector of \( D g_{ss}(s_t) \) and \( D_j h_{ss}(s_t) \) is the \( j^{th} \) column vector of \( D h_{ss}(s_t) \). To simply notation, define
\[
D_{n,m} g_{ss}(s_t) \equiv [D_n g_{ss}(s_t), \ldots, D_m g_{ss}(s_t)] \quad \text{and} \quad D_{n,m} h_{ss}(s_t) \equiv [D_n h_{ss}(s_t), \ldots, D_m h_{ss}(s_t)]
\]
for all \( n \) and \( m \) and all \( s_t \).

Hence, the above approximations are equivalent to
\[
g(x_{t-1}, \varepsilon_t, \chi, s_t) - y_{ss} \simeq D_{1,n_x} g_{ss}(s_t) (x_{t-1} - x_{ss}) + D_{n_x+1,n_x+n_s} g_{ss}(s_t) \varepsilon_t + D_{n_x+n_s+1} g_{ss}(s_t) \chi
\]
and
\[
h(x_{t-1}, \varepsilon_t, \chi, s_t) - x_{ss} \simeq D_{1,n_x} h_{ss}(s_t) (x_{t-1} - x_{ss}) + D_{n_x+1,n_x+n_s} h_{ss}(s_t) \varepsilon_t + D_{n_x+n_s+1} h_{ss}(s_t) \chi
\]

The objective is now to find the coefficients
\[
\{D_{1,n_x} g_{ss}(s_t), D_{1,n_x} h_{ss}(s_t)\}_{s_t=1}^{n_s} \quad \{D_{n_x+1,n_x+n_s} g_{ss}(s_t), D_{n_x+1,n_x+n_s} h_{ss}(s_t)\}_{s_t=1}^{n_s} \quad \text{and} \quad \{D_{n_x+n_s+1} g_{ss}(s_t), D_{n_x+n_s+1} h_{ss}(s_t)\}_{s_t=1}^{n_s}
\]
of the above describe expansions, where $D_{1,n_x}g_{ss} (s_t) \in \mathbb{C}^{n_y \times n_x}$, $D_{1,n_x}h_{ss} (s_t) \in \mathbb{C}^{n_x \times n_x}$, $D_{n_x+1,n_x+n_c}g_{ss} (s_t) \in \mathbb{C}^{n_y \times n_c}$, $D_{n_x+1,n_x+n_c}h_{ss} (s_t) \in \mathbb{C}^{n_x \times n_c}$, $D_{n_x+n_c+1}g_{ss} (s_t) \in \mathbb{C}^{n_y \times 1}$, and $D_{n_x+n_c+1}h_{ss} (s_t) \in \mathbb{C}^{n_x \times 1}$

for all $s_t$ and where we are using $\mathbb{C}^{m_1 \times m_2}$ to denote $m_1 \times m_2$ matrices over the complex numbers.

The current setup requires finding a set of $n_s$ policy functions, one for each possible value of the Markov chain, instead of the single set of policy functions in the constant parameter case.

The coefficients of these policy functions are going to be obtained by using the fact that

$$G(\bar{x}_t; t, s_t) = 0$$

for all $\bar{x}_t$; $t$; $s_t$ and, therefore, it must be the case that

$$DG(\bar{x}_t; t, s_t) = 0$$

for all $\bar{x}_t$, $t$, and $s_t$ and, in particular,

$$DG_{ss} (s_t) = 0$$

for all $s_t$. Thus,

$$[D_1G_{ss} (s_t), \ldots, D_{n_x}G_{ss} (s_t)] = 0_{(n_y+n_x)\times n_x},$$

$$[D_{n_x+1}G_{ss} (s_t), \ldots, D_{n_x+n_c}G_{ss} (s_t)] = 0_{(n_y+n_x)\times n_c},$$

and $D_{n_x+n_c+1}G_{ss} (s_t) = 0_{n_y+n_x}$

for all $s_t$ where $D_jG_{ss} (s_t)$ is the $j^{th}$ column vector of $DG_{ss} (s_t)$. Again, note that there are a set of $n_s$ derivatives of $G$, one for each possible value of $s_t$, instead of the single derivative in the constant parameter case. To simply notation, again, define

$$D_{n,n_c}G_{ss} (s_t) \equiv [D_nG_{ss} (s_t), \ldots, D_{n}G_{ss} (s_t)]$$

for all $s_t$. Therefore, expression (10) can be written as

$$D_{1,n_x}G_{ss} (s_t) = 0_{(n_y+n_x)\times n_x}, \quad D_{n_x+1,n_x+n_c}G_{ss} (s_t) = 0_{(n_y+n_x)\times n_c}, \quad \text{and} \quad D_{n_x+n_c+1}G_{ss} (s_t) = 0_{n_y+n_x}$$

(11)

for all $s_t$. 

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The expression $\mathcal{D}_{1,n_x} \mathbb{G}_{ss} (s_t) = 0_{(n_y+n_z) \times n_z}$ will be used to solve for $\{ \mathcal{D}_{1,n_x} g_{ss} (s_t) , \mathcal{D}_{1,n_x} h_{ss} (s_t) \}_{s_t=1}^{n_s}$, while expressions $\mathcal{D}_{n_x+1,n_x+n_y} \mathbb{G}_{ss} (s_t) = 0_{(n_y+n_z) \times n_z}$ and $\mathcal{D}_{n_x+n_y+1} \mathbb{G}_{ss} (s_t) = 0_{n_y+n_z}$ will be used to solve for $\{ \mathcal{D}_{n_x+1,n_x+n_y} g_{ss} (s_t) , \mathcal{D}_{n_x+1,n_x+n_y} h_{ss} (s_t) \}_{s_t=1}^{n_s}$ and $\{ \mathcal{D}_{n_x+n_y+1} g_{ss} (s_t) , \mathcal{D}_{n_x+n_y+1} h_{ss} (s_t) \}_{s_t=1}^{n_s}$.

In what follows, we show that $\mathcal{D}_{1,n_x} \mathbb{G}_{ss} (s_t) = 0_{(n_y+n_z) \times n_z}$ implies a quadratic system, while $\mathcal{D}_{n_x+1,n_x+n_y} \mathbb{G}_{ss} (s_t) = 0_{(n_y+n_z) \times n_z}$ and $\mathcal{D}_{n_x+n_y+1} \mathbb{G}_{ss} (s_t) = 0_{n_y+n_z}$ (given $\{ \mathcal{D}_{1,n_x} g_{ss} (s_t) , \mathcal{D}_{1,n_x} h_{ss} (s_t) \}_{s_t=1}^{n_s}$) are linear systems.

### 3.1 Solving for the Derivatives of $x_{t-1}$

Taking derivatives with respect to $x_{t-1}$ in (9) produces the following expression

$$\mathcal{D}_{1,n_x} \mathbb{G}_{ss} (s_t) = \sum_{s'=1}^{n_s} p_{s_t,s'} \int \left( \begin{array}{l} \mathcal{D}_{1,n_y} f_{ss} (s', s_t) \mathcal{D}_{1,n_x} g_{ss} (s') \mathcal{D}_{1,n_x} h_{ss} (s_t) \\ + \mathcal{D}_{n_y+1,2n_y} f_{ss} (s', s_t) \mathcal{D}_{1,n_x} g_{ss} (s_t) \\ + \mathcal{D}_{2n_y+1,2n_y+n_z} f_{ss} (s', s_t) \mathcal{D}_{1,n_x} h_{ss} (s_t) \\ + \mathcal{D}_{2n_y+n_x+1,2(n_y+n_z)} f_{ss} (s', s_t) \end{array} \right) \mu (\varepsilon') d\varepsilon'$$

for all $s_t$. Next, taking into account that $\int \mu (\varepsilon') d\varepsilon' = 1$, this expression simplifies to

$$\mathcal{D}_{1,n_x} \mathbb{G}_{ss} (s_t) = \sum_{s'=1}^{n_s} p_{s_t,s'} \left( \begin{array}{l} \mathcal{D}_{1,n_y} f_{ss} (s', s_t) \mathcal{D}_{1,n_x} g_{ss} (s') \mathcal{D}_{1,n_x} h_{ss} (s_t) \\ + \mathcal{D}_{n_y+1,2n_y} f_{ss} (s', s_t) \mathcal{D}_{1,n_x} g_{ss} (s_t) \\ + \mathcal{D}_{2n_y+1,2n_y+n_z} f_{ss} (s', s_t) \mathcal{D}_{1,n_x} h_{ss} (s_t) \\ + \mathcal{D}_{2n_y+n_x+1,2(n_y+n_z)} f_{ss} (s', s_t) \end{array} \right)$$

for all $s_t$. Now, rearranging, for each $s_t$

$$\mathcal{D}_{1,n_x} \mathbb{G}_{ss} (s_t) = \sum_{s'=1}^{n_s} p_{s_t,s'} \left( \begin{array}{l} \mathcal{D}_{1,n_y} f_{ss} (s', s_t) \mathcal{D}_{1,n_x} g_{ss} (s') + \mathcal{D}_{n_y+1,2n_y+n_z} f_{ss} (s', s_t) \mathcal{D}_{1,n_x} h_{ss} (s_t) \\ + \mathcal{D}_{n_y+1,2n_y} f_{ss} (s', s_t) \mathcal{D}_{1,n_x} g_{ss} (s_t) + \mathcal{D}_{2n_y+n_x+1,2(n_y+n_z)} f_{ss} (s', s_t) \end{array} \right).$$

Putting together the $n_s$ versions of (12), one for each value of $s_t$, and equating them to zero, as implied by (11), yields a system of $(n_y + n_z) n_x n_s$ quadratic equations in the same number of unknowns $\{ \mathcal{D}_{1,n_x} g_{ss} (s_t) , \mathcal{D}_{1,n_x} h_{ss} (s_t) \}_{s_t=1}^{n_s}$. Section 4 describes how to solve this system.
3.2 Solving for the Derivatives of $\varepsilon_t$ and $\chi$

This subsection shows that, after finding \( \{D_{n_x+n_z} g_{ss} (s_t), D_{n_x+n_z} h_{ss} (s_t)\} \) for all \( s_t \), obtaining \( \{D_{n_x+1, n_x+n_z} g_{ss} (s_t), D_{n_x+1, n_x+n_z} h_{ss} (s_t)\} \) and \( \{D_{n_x+n_z+1} g_{ss} (s_t), D_{n_x+n_z+1} h_{ss} (s_t)\} \) is simply solving a system of linear equations. The first step is to solve for \( \{D_{n_x+1, n_x+n_z} g_{ss} (s_t), D_{n_x+1, n_x+n_z} h_{ss} (s_t)\} \). Then, we solve for \( \{D_{n_x+n_z+1} g_{ss} (s_t), D_{n_x+n_z+1} h_{ss} (s_t)\} \).

3.2.1 Solving for the Derivatives of $\varepsilon_t$

In order to solve for \( \{D_{n_x+1, n_x+n_z} g_{ss} (s_t), D_{n_x+1, n_x+n_z} h_{ss} (s_t)\} \), we obtain the expression for \( D_{n_x+1, n_x+n_z} g_{ss} (s_t) \) by taking derivatives with respect to $\varepsilon_t$ in (9)

\[
D_{n_x+1, n_x+n_z} g_{ss} (s_t) =
\sum_{s'=1}^{n_s} p_{s_t,s'} \int \left( D_{1,n_y} f_{ss} (s', s_t) D_{n_x+n_z} g_{ss} (s') D_{n_x+1,n_x+n_z} h_{ss} (s_t) + D_{n_y+2,n_y} f_{ss} (s', s_t) D_{n_x+n_z} g_{ss} (s_t) + D_{2,n_y+2,n_y+n_z} f_{ss} (s', s_t) D_{n_x+n_z} h_{ss} (s_t) + D_{2,n_y+n_z+1,2(n_y+n_z+n_z)} f_{ss} (s', s_t) \right) \mu (\varepsilon') d\varepsilon'
\]

for all \( s_t \). Taking into account that \( \int \mu (\varepsilon') d\varepsilon' = 1 \), this expression simplifies to

\[
D_{n_x+1, n_x+n_z} g_{ss} (s_t) =
\sum_{s'=1}^{n_s} p_{s_t,s'} \left( D_{1,n_y} f_{ss} (s', s_t) D_{n_x+n_z} g_{ss} (s') + D_{n_y+2,n_y} f_{ss} (s', s_t) D_{n_x+n_z} g_{ss} (s_t) + D_{2,n_y+n_z+1,2(n_y+n_z+n_z)} f_{ss} (s', s_t) \right) D_{n_x+n_z+1} h_{ss} (s_t)
\]

(13)

for all \( s_t \).

Putting together the \( n_s \) versions of (13), one for each value of \( s_t \), and equating them to zero, as implied by (11), yields a system of \( (n_y + n_x) n_\varepsilon n_s \) equations in the same number of unknowns \( \{D_{n_x+1, n_x+n_z} g_{ss} (s_t), D_{n_x+1, n_x+n_z} h_{ss} (s_t)\} \). The system is linear.
The linear system can be written in matrix notation expression as

\[
\begin{bmatrix}
\Theta_\varepsilon & \Phi_\varepsilon
\end{bmatrix}
\begin{bmatrix}
D_{n_x+1,n_x+n_x+gss} (1) \\
\vdots \\
D_{n_x+1,n_x+n_x+gss} (n_s) \\
D_{n_x+1,n_x+n_x+hss} (1) \\
\vdots \\
D_{n_x+1,n_x+n_x+hss} (n_s)
\end{bmatrix} = \Psi_\varepsilon
\]  

(14)

where

\[
\Theta_\varepsilon = \sum_{s'=1}^{n_s} \begin{bmatrix}
p_{1,s'} D_{n_y+1,2n_y+fs} (s', 1) & \cdots & 0_{(n_x+n_y)\times n_y} \\
\vdots & \ddots & \vdots \\
0_{(n_x+n_y)\times n_y} & \cdots & p_{n_s,s'} D_{n_y+1,2n_y+fs} (s', n_s)
\end{bmatrix},
\]

\[
\Phi_\varepsilon = \sum_{s'=1}^{n_s} \begin{bmatrix}
p_{1,s'} D_{1,n_x+fs} (s', 1) & \cdots & 0_{(n_x+n_y)\times n_x} \\
\vdots & \ddots & \vdots \\
0_{(n_x+n_y)\times n_x} & \cdots & p_{n_s,s'} D_{1,n_x+fs} (s', n_s) D_{1,n_x+gss} (s')
\end{bmatrix}
\]

and

\[
\Psi_\varepsilon = -\sum_{s'=1}^{n_s} \begin{bmatrix}
p_{1,s'} D_{2(n_y+n_x)+n_x+1,2(n_y+n_x+n_x)+fs} (s', 1) \\
\vdots \\
p_{n_s,s'} D_{2(n_y+n_x)+n_x+1,2(n_y+n_x+n_x)+fs} (s', n_s)
\end{bmatrix}.
\]

Thus, given the solution for \( \{D_{1,n_x+gss} (s_t), D_{1,n_x+hss} (s_t)\}_{s_t=1}^{n_s} \), expression (14) is a system of \((n_y+n_x)n_xn_x\) linear equations in the same number of unknowns given by \( \{D_{n_x+1,n_x+n_x+gss} (s_t), D_{n_x+1,n_x+n_x+hss} (s_t)\}_{s_t=1}^{n_s} \) that can be solved by inverting \( [\Theta_\varepsilon \Phi_\varepsilon] \).
3.2.2 Solving for the Derivatives of $\chi$

In order to solve for $\{D_{n_x+n_c+1}g_{ss} (s_t), D_{n_x+n_c+1}h_{ss} (s_t)\}_{s_t=1}^{n_s}$, we obtain the expression for $D_{n_x+n_c+1}G_{ss} (s_t)$ by taking derivatives with respect to $\chi$ in (9)

$$D_{n_x+n_c+1}G_{ss} (s_t) =$$

$$\sum_{s' = 1}^{n_s} p_{s_t, s'} \int \left[ D_{1, n_y} f_{ss} (s', s_t) \left[ D_{1, n_y} g_{ss} (s') D_{n_x+n_c+1}h_{ss} (s_t) + D_{n_x+n_c+1}g_{ss} (s') \varepsilon' + D_{n_x+n_c+1}h_{ss} (s') \right] + \mu (\varepsilon') d\varepsilon' \right]$$

for all $s_t$, where $D\theta_{ss} (s_t)$ is the derivative of $\theta (\chi, s_t)$ with respect to $\chi$ evaluated at $\chi = 0$

$$D\theta_{ss} (s_t) = D\theta (0, s_t) = [D_j^i \theta (0, s_t)]_{1 \leq i \leq n_g, j = 1}$$

for all $s_t$.

Taking into account that $\int \mu (\varepsilon') d\varepsilon' = 1$ and $\int \varepsilon' \mu (\varepsilon') d\varepsilon' = 0$, the above simplifies to

$$D_{n_x+n_c+1}G_{ss} (s_t) =$$

$$\sum_{s' = 1}^{n_s} p_{s_t, s'} \left[ D_{n_y+1, 2n_y} f_{ss} (s', s_t) \left\{ D_{1, n_y} g_{ss} (s') D_{n_x+n_c+1}h_{ss} (s_t) + D_{n_x+n_c+1}g_{ss} (s') \right\} + D_{n_y+n_x+n_c+1}g_{ss} (s_t) \right]$$

$$+ D_{1, n_y} f_{ss} (s', s_t) \left[ D_{n_y+1, 2n_y} g_{ss} (s') D_{n_x+n_c+1}h_{ss} (s_t) + D_{2, n_y+1, 2n_y+n_x} f_{ss} (s', s_t) D_{n_x+n_c+1}h_{ss} (s_t) \right]$$

$$+ D_{2, n_y+n_x+n_c+1, 2(n_y+n_x+n_c)+n_y} f_{ss} (s', s_t) \left\{ D\theta_{ss} (s') + D_{1, n_y} g_{ss} (s') D_{n_x+n_c+1}h_{ss} (s_t) \right\}$$

for all $s_t$.

Putting together the $n_s$ versions of (15), one for each value of $s_t$, and equating them to zero, as implied by (11), yields a system of $(n_y + n_x) n_s$ equations in the same number of unknowns $\{D_{n_x+n_c+1}g_{ss} (s_t), D_{n_x+n_c+1}h_{ss} (s_t)\}_{s_t=1}^{n_s}$. This system is also linear.
The linear system can be written in matrix notation expression as

\[
\begin{bmatrix}
\Theta \chi & \Phi \chi \\
\end{bmatrix}
\begin{bmatrix}
\mathcal{D}_{n_x+n_c+1} s \ss (1) \\
\vdots \\
\mathcal{D}_{n_x+n_c+1} s \ss (n_x) \\
\mathcal{D}_{n_x+n_c+1} h \ss (1) \\
\vdots \\
\mathcal{D}_{n_x+n_c+1} h \ss (n_x) \\
\end{bmatrix} = \Psi \chi,
\]

where

\[
\Theta \chi = \sum_{s'=1}^{n_x} \begin{bmatrix}
p_{1,s'} \mathcal{D}_{n_y+1} s \ss (s', 1) & \cdots & 0_{(n_x+n_y) \times n_y} \\
\vdots & \ddots & \vdots \\
0_{(n_x+n_y) \times n_y} & \cdots & p_{n_x,s'} \mathcal{D}_{n_y+1} s \ss (s', n_x) \\
\end{bmatrix}
\]

and

\[
\Phi \chi = \sum_{s'=1}^{n_x} \begin{bmatrix}
p_{1,s'} \mathcal{D}_{1} s \ss (s', 1) \mathcal{D}_{1} s \ss (s') & \cdots & 0_{(n_x+n_y) \times n_x} \\
\vdots & \ddots & \vdots \\
0_{(n_x+n_y) \times n_x} & \cdots & p_{n_x,s'} \mathcal{D}_{1} s \ss (s', n_x) \mathcal{D}_{1} s \ss (s') \\
\end{bmatrix}
\]

\[
\Psi \chi = - \sum_{s'=1}^{n_x} \begin{bmatrix}
p_{1,s'} \left( \mathcal{D}_{2} (n_y+n_x+n_c+n_\theta+1) s \ss (s', 1) \mathcal{D}_{s \ss} (s') + \cdots \right) \\
\vdots \\
p_{n_x,s'} \left( \mathcal{D}_{2} (n_y+n_x+n_c+n_\theta+1) s \ss (s', n_x) \mathcal{D}_{s \ss} (s') + \cdots \right) \\
\end{bmatrix}.
\]
Thus, given the solution for \( \{D_{1,n_x}g_{ss}(s_t), D_{1,n_x}h_{ss}(s_t)\}_{s_t=1}^{n_s} \), expression (16) is a system of \((n_y + n_x)n_s\) linear equations in the same number of unknowns given by \( \{D_{n_x+n_y+1}g_{ss}(s_t), D_{n_x+n_y+1}h_{ss}(s_t)\}_{s_t=1}^{n_s} \) that can be solved by inverting \( \left[ \Theta_X \Phi_X \right] \).

### 3.3 Non-Certainty Equivalence of First-Order Approximation

As pointed out by ?, one important feature of constant parameter models is certainty equivalence of the first-order approximation. This feature of constant parameter models implies that first-order approximations are inadequate for analyzing interesting behavior such as responses to risk because the approximated decision rules are invariant to changes in volatility. For example, ? and ? note that at least second-order approximations are needed to analyze certain asset pricing implications, such as the yield curve, since second-order approximations are not certainty equivalent, and hence react to changes in volatility. Second-order approximations also imply a degree of difficulty in performing likelihood based estimation, such as ? who use the particle filter for estimation. These factors mean that addressing interesting questions with second-order approximations may be necessary but difficult in constant parameter models. As shown below, first order approximations to Markov Switching models are not necessarily certainty equivalent. This nice feature opens the door to analyze risk related behaviors using linearly approximated models.

To see the certainty equivalence of the first-order approximation of constant parameter models, consider equation (16) with only one regime, so \( n_s = 1 \). In this case

\[
\left[ \begin{array}{c} \Theta_X \\ \Phi_X \end{array} \right] \left[ \begin{array}{c} D_{n_x+n_y+1}g_{ss}(1) \\ D_{n_x+n_y+1}h_{ss}(1) \end{array} \right] = \Psi_X
\]

where

\[
\left[ \begin{array}{c} \Theta_X \\ \Phi_X \end{array} \right] =
\left[ D_{1,n_y}f_{ss}(1,1)D_{1,n_x}g_{ss}(1) + D_{2,n_y+1,2n_y+n_x}f_{ss}(1,1) \right] D_{1,n_y}f_{ss}(1,1) + D_{n_y+1,2n_y}f_{ss}(1,1)
\]

and

\[
\Psi_X = - \left[ D_{2(n_y+n_x+n_c)+1,2(n_y+n_x+n_c)+n_y}f_{ss}(1,1)D_{ss} \theta(1) \right].
\]
Clearly, in constant parameter models it is the case that \( \theta(\chi, 1) = \bar{\theta} \). Therefore, \( D_{ss}\theta(1) = 0_{n_y} \), which implies \( \Psi_\chi = 0_{n_x+n_y} \). Consequently the system (17) is homogenous. If a unique solution exists, then it is given by

\[
D_{n_x+n_x+1g_{ss}}(1) = 0_{n_y} \text{ and } D_{n_x+n_x+1h_{ss}}(1) = 0_{n_x}.
\] (18)

Remenber that in constant parameter models the linear approximation to the policy rules imply

\[
y_t - y_{ss} = D_{1,n_xg_{ss}}(1)(x_{t-1} - x_{ss}) + D_{n_x+n_x+n_xg_{ss}}(1)\varepsilon_t + D_{n_x+n_x+1g_{ss}}(1)
\]
and

\[
x_t - x_{ss} = D_{1,n_xh_{ss}}(1)(x_{t-1} - x_{ss}) + D_{n_x+n_x+n_xh_{ss}}(1)\varepsilon_t + D_{n_x+n_x+1h_{ss}}(1).
\]

Using (18), evaluated at \( x_{t-1} = x_{ss} \) and \( \varepsilon_t = 0_{n_x} \), the above approximations imply that \( y_t - y_{ss} = 0_{n_y} \) and \( x_t - x_{ss} = 0_{n_x} \), i.e. the linear approximation of constant parameter models is certain equivalent.

Let us now turn to the Markov switching case. From equation (16) is clear that a necessary condition for the linear approximation not to be certain equivalent is that \( \Psi_\chi \neq 0_{n_x(n_x+n_y)} \). Let us analize when it is the case that \( \Psi_\chi \neq 0_{n_x(n_x+n_y)} \). Consider the expression for \( \Psi_\chi \)

\[
\Psi_\chi = -\sum_{s' = 1}^{n_x} p_{s,s'} \begin{bmatrix}
D_{2(n_y+n_x+n_x)+1,2(n_y+n_x+n_x)+n_y}f_{ss}(s', 1)\mathcal{D}_{\theta_{ss}}(s') + \ldots \\
D_{2(n_y+n_x+n_x)+n_y+1,2(n_y+n_x+n_x)+n_y}f_{ss}(s', 1)\mathcal{D}_{\theta_{ss}}(1) \\
\vdots \\
p_{n_x,s'} \begin{bmatrix}
D_{2(n_y+n_x+n_x)+1,2(n_y+n_x+n_x)+n_y}f_{ss}(s', n_x)\mathcal{D}_{\theta_{ss}}(s') + \ldots \\
D_{2(n_y+n_x+n_x)+n_y+1,2(n_y+n_x+n_x)+n_y}f_{ss}(s', n_x)\mathcal{D}_{\theta_{ss}}(n_x)
\end{bmatrix}
\end{bmatrix}.
\]

Then, if \( \mathcal{D}_{\theta_{ss}}(s_t) = 0_{n_y} \) for all \( s_t \), it is the case that \( \Psi_\chi = 0_{n_x(n_x+n_y)} \). So a necessary condition for \( \Psi_\chi \neq 0_{n_x(n_x+n_y)} \) is that \( \mathcal{D}_{\theta_{ss}}(s_t) \neq 0_{n_y} \) for some \( s_t \). Recalling the form of \( \theta_t \)

\[
\theta_1(\chi, s_t) = \overline{\theta}_1 + \chi \tilde{\theta}_1(s_t)
\]
and

\[
\theta_2(\chi, s_t) = \tilde{\theta}_2(s_t)
\]
we conclude that

$$D\theta_{ss}(s_t) = \left[ \hat{\theta}_1(s_t) 0 \right]'$$.

Then $D\theta_{ss}(s_t) \neq 0_n$ for some $s_t$ if and only if $\hat{\theta}_1(s_t) \neq 0_n$ for some $s_t$. Hence, a necessary condition for $\Psi_{\lambda} \neq 0_{n_s(n_x+n_y)}$ is that $\hat{\theta}_1(s_t) \neq 0_n$ for some $s_t$.

However, the condition that $\hat{\theta}_1(s_t) \neq 0_n$ for some $s_t$ is not sufficient for $\Psi_{\lambda} \neq 0_{n_s(n_x+n_y)}$. In addition, it must be the case that

$$\sum_{s'=1}^{n_s} p_{st,s'} \left( D_{2(n_y+n_x+n_x)+1,2(n_y+n_x+n_x)+n_y} f_{ss}(s',s_t) D\theta_{ss}(s') + D_{2(n_y+n_x+n_x)+n_y+1,2(n_y+n_x+n_x)+n_y} f_{ss}(s',s_t) D\theta_{ss}(s_t) \right) \neq 0_{n_x+n_y}$$

for some $s_t$, which will be true when $\theta_{1t}$ do not enter the equilibrium conditions multiplicatively with a variable which expected value equals zero when evaluating $f(\cdot)$ in steady state. The following Proposition summarizes these results.

**Proposition 1** Let $\hat{\theta}_1(s_t) = 0_n$ for all $s_t$. Then $\Psi_{\lambda} = 0_{n_s(n_x+n_y)}$ and

$$D_{n_x+n_x+1} s_{ss}(s_t) = 0_{n_y}$$

and

$$D_{n_x+n_x+1} h_{ss}(s_t) = 0_{n_x}$$

for all $s_t$ and the first order approximation is certainty equivalent. On the other hand, let $\hat{\theta}_1(s_t) \neq 0_n$ for some $s_t$ but let

$$\sum_{s'=1}^{n_s} p_{st,s'} \left( D_{2(n_y+n_x+n_x)+1,2(n_y+n_x+n_x)+n_y} f_{ss}(s',s_t) D\theta_{ss}(s') + D_{2(n_y+n_x+n_x)+n_y+1,2(n_y+n_x+n_x)+n_y} f_{ss}(s',s_t) D\theta_{ss}(s_t) \right) = 0_{n_x+n_y}$$

for all $s_t$. Then $\Psi_{\lambda} = 0_{n_s(n_x+n_y)}$ and

$$D_{n_x+n_x+1} s_{ss}(s_t) = 0_{n_y}$$

and

$$D_{n_x+n_x+1} h_{ss}(s_t) = 0_{n_x}$$

for all $s_t$ and the first order approximation is certainty equivalent.

Note that if the system is not certainty equivalent, it means that either

$$D_{n_x+n_x+1} s_{ss}(s_t) \neq 0_{n_y}$$

or

$$D_{n_x+n_x+1} h_{ss}(s_t) \neq 0_{n_x}$$
for some $s_t$ and the linear approximation to the policy rules evaluated at $x_{t-1} = x_{ss}$ and $\varepsilon_t = 0$ imply that either
\[
y_t - y_{ss} = \mathcal{D}_{n_y+n_x+1}g_{ss} (s_t) \neq 0_{n_y}
\]
or
\[
x_t - x_{ss} = \mathcal{D}_{n_y+n_x+1}h_{ss} (s_t) \neq 0_{n_x}
\]
for some $s_t$.

### 4 The Solution to the Quadratic System

As mentioned above, the $n_s$ versions of (12) form a system of $(n_y + n_x) n_s n_x$ quadratic equations in the elements of $\{\mathcal{D}_{1,n_x}g_{ss} (s), \mathcal{D}_{1,n_x}h_{ss} (s)\}_{s=1}^{n_s}$. This section describes how to find the solution to this system. Putting (12) into matrix form produces $n_s$ systems of the form

\[
A (s_t) \begin{bmatrix} I \\ \mathcal{D}_{1,n_x}g_{ss} (1) \\ \vdots \\ \mathcal{D}_{1,n_x}g_{ss} (n_s) \end{bmatrix} \mathcal{D}_{1,n_x}h_{ss} (s_t) = B (s_t) \begin{bmatrix} I \\ \mathcal{D}_{1,n_x}g_{ss} (s_t) \end{bmatrix}
\]

(19)

for all $s_t$, where

\[
A (s_t) = \left[ \sum_{s'=1}^{n_s} p_{s_t,s'} \mathcal{D}_{2n_y+n_x+1,2n_y+n_x}f_{ss} (s', s_t) \right. p_{s_t,1} \mathcal{D}_{1,n_y}f_{ss} (1, s_t) \left. \cdots p_{s_t,n_x} \mathcal{D}_{1,n_y}f_{ss} (n_x, s_t) \right]
\]

and

\[
B (s_t) = -\sum_{s'=1}^{n_s} p_{s_t,s'} \left[ \mathcal{D}_{2n_y+n_x+1,2(n_y+n_x)}f_{ss} (s', s_t) \mathcal{D}_{n_y+1,2n_y}f_{ss} (s', s_t) \right].
\]

This quadratic system is nothing else than an algebraic system of equations. In a constant regime framework, $n_s = 1$, mapping this system into a generalized eigenvalue problem allows solving it by a singular value decomposition (SVD) type of algorithm. In the case of Markov switching, the fact that $\{\mathcal{D}_{1,n_x}g_{ss} (s)\}_{s=1}^{n_s}$ appear in every of the $n_s$ equations described above makes it impossible to map the algebraic systems of equations into a generalized eigenvalue problem. Instead solutions are found using Gröbner Bases.
4.1 Gröbner Basis

What is a Gröbner basis? A Gröbner basis for a system of polynomials is a set of multivariate polynomials that possesses desirable algorithmic properties. The most important of these features for the current problem is that the system of polynomials in a Gröbner basis have the same collection of roots as the original polynomials. Every set of polynomials can be transformed into a Gröbner basis, although this transformation may not be unique. The transformation process generalizes the familiar techniques of Gaussian elimination for solving linear systems of equations. In general, solving the problem in the system of polynomials in a Gröbner basis is much simpler than in the original system. Also, a fundamental insight and contribution of Gröbner bases theory is that every polynomial system, no matter how complicated, can be transformed into Gröbner basis form (see Buchberger’s algorithm).

As an example, consider the following system of polynomials of four quadratic equations in four unknowns:

\[
\begin{align*}
xy + zw + 2 & = 0, \\
xy + yz + 3 & = 0, \\
xz + wx + wy + 6 & = 0, \text{ and} \\
xz + 2xy + 3 & = 0.
\end{align*}
\]

A Gröbner basis, with respect to the lexicographic ordering \(\{x, y, z, w\}\), is

\[
\begin{align*}
-49 - 19w^2 + 9w^4 + 3w^6, \\
2w + 9w^3 + 3w^5 + 14z, \\
-99w + 6w^3 + 9w^5 + 28y, \text{ and} \\
15w - 6w^3 - 9w^5 + 28x.
\end{align*}
\]

Note that the first element of the basis is a polynomial in \(w\) only. Given a root \(w\) of the first polynomial, the second polynomial is linear in \(z\), the third is linear in \(y\), and the last is linear in \(x\). Solving the first element of the basis produces the following six solutions

\(\{w = -1.55461, w = -1.39592i, w = 1.39592i, w = 0, -1.86232i, w = 1.86232i, w = 1.55461\}\)
Solving the other three basis, conditional on these solutions, gives the following roots

\[\{z = 4.58328, z = -0.41342i, z = 0.41342i, z = 0.914097i, z = 0.914097i, z = -4.58328\}\],

\[\{y = -1.7728, y = -3.81477i, y = 3.81477i, y = -0.768342i, y = 0.768342i, y = 1.7728\}\]

and

\[\{x = -2.89104, x = -0.372997i, x = 0.372997i, x = -4.81861i, x = 4.81861i, x = 2.89104\}\].

These roots solve the original system of four quadratic equations in four unknowns.

### 4.2 Number of Solutions

In the previous example, there are six solutions to the set of quadratic equations. This result follows from the fact that the first polynomial in the Gröbner basis is a sixth order polynomial in \(w\) and the following three equations are linear in \(x\), \(y\), and \(z\). In general, determining the number of solutions from an arbitrary original set of polynomials may be difficult. However, given the structure of the quadratic system (19), it is possible to characterize the number of solutions.

Consider the fixed regime case of equation (19), when \(n_s = 1\). The usual practice of solving the model, as in ?, involves constructing a single stable solution that depends upon the generalized eigenvalues of the matrices \(A(1)\) and \(B(1)\). The full set of solutions (stable and unstable) can be found by different selections of eigenvalues. The total combination of eigenvalues depends upon the rank of the matrix \(A(1)\) and the number of exogenous predetermined variables by the following.

**Proposition 2 Fixed Regime Case.** Let \(n_{exo}\) denote the number of exogenous predetermined variables, so \(0 \leq n_{exo} \leq n_x\). Then the total number of solutions to equation (19) when \(n_s = 1\) is given by

\[
\binom{\text{rank}A(1) - n_{exo}}{n_x - n_{exo}} = \frac{(\text{rank}A(1) - n_{exo})!}{(n_x - n_{exo})!(\text{rank}A(1) - n_x)!}.
\]
The matrix $A$ will not be of full rank when there exist redundant variables or identities that can be eliminated from the system, and consequently are linearly dependent upon another set of variables.\(^3\) The set of exogenous variables has set eigenvalues associated with their coefficients of autocorrelation, and consequently the number of solutions doesn’t depend upon the number of exogenous variables.

Now consider the case with $n_s > 1$. Given the previous proposition, the number of solutions to each of the $n_s$ equations of (19) depends on the rank of $A(s_t)$ and the number of exogenous variables. Since there are $n_s$ quadratic equations of this form, the total number of solutions depends on all possible combinations of solutions of the form in the previous proposition.

**Proposition 3** Switching Case. Let $n_{exo}$ denote the number of exogenous predetermined variables, so $0 \leq n_{exo} \leq n_x$. Then the total number of solutions to the $n_s$ equations (19) is given by

$$
\prod_{s=1}^{n_s} \left( \frac{\text{rank} A(s) - n_{exo}}{n_x - n_{exo}} \right) = \prod_{s=1}^{n_s} \left( \frac{(\text{rank} A(s) - n_{exo})!}{(n_x - n_{exo})! (\text{rank} A(s) - n_x)!} \right)
$$

Consequently, the Gröbner basis method will return all the possible solutions to the quadratic system (19). Given the full set of possible solutions, now it is possible to use a definition of stability to determine how many solutions are stable.

### 4.3 Mean Square Stability

The Gröbner basis methodology will return all possible solutions of the $n_s$ systems (19) in the unknowns $\{D_{1,n_s} h_{ss}(s), D_{1,n_s} g_{ss}(s)\}_{s=1}^{n_s}$. The issue now is whether any of the solutions are stable, and if so, how many.

In a typical model without Markov switching, determinacy is easily verified by checking whether the number of eigenvalues of the system (19) inside the unit circle equals to the number of state variables. In a model with Markov switching, as the one described here, the problem

\(^3\)A simple example of a system with $A(1)$ of nonfull rank is the RBC model with capital, consumption, and output. Often the output variable is eliminated, since it is redundant given a solution for consumption and capital.
is more subtle. As shown in ?, it is possible that the number of stable eigenvalues associated
with each of the regimes is equal to the number of states but the system, as a whole, does not
have a stable solution under several concepts of stability. The good news is that the Markov
switching model can be checked for mean-square stability (MSS), as de…ned in ?. In particular,
MSS requires checking if the following matrix has its eigenvalues inside the unit circle
\[ T = (P' \otimes I_{n_z^2}) \text{diag} [\mathcal{D}_{1,n_x} h_{ss}(s) \otimes \mathcal{D}_{1,n_x} h_{ss}(s)] \] (20)
where
\[
\text{diag} [\mathcal{D}_{1,n_x} h_{ss}(s) \otimes \mathcal{D}_{1,n_x} h_{ss}(s)] =
\begin{bmatrix}
\mathcal{D}_{1,n_x} h_{ss}(1) \otimes \mathcal{D}_{1,n_x} h_{ss}(1) & 0_{n_x^2 \times n_z^2} & \cdots \\
0_{n_x^2 \times n_z^2} & \mathcal{D}_{1,n_x} h_{ss}(2) \otimes \mathcal{D}_{1,n_x} h_{ss}(2) & \cdots \\
\vdots & \vdots & \ddots \\
0_{n_x^2 \times n_z^2} & 0_{n_x^2 \times n_z^2} & \mathcal{D}_{1,n_x} h_{ss}(n_z) 
\end{bmatrix}
\]
Thus, with Markov switching, the policy functions \{\mathcal{D}_{1,n_x} h_{ss}(i)\}_{i=1}^{n_z} for all possible solutions
must be checked for stability under (20). If only one policy function is stable then the model only
has one stable solution. If more than one are stable, the model has multiple stable solutions. If
none are stable, the model has no stable solutions.

5 Second Order Approximation

Having constructed the …rst-order approximations, this Section shows how to …nd the second
order Taylor expansions to \( g \) and \( h \) around the point \((x_{ss}, 0_{n_x}, 0, s_t)\). These second order
expansions have the form, for each \( i = 1, \ldots, n_y \):

\[
g^i (x_{t-1}, \epsilon_t, \chi, s_t) - y_{ss} \simeq \mathcal{D}g^i_{ss}(s_t) \begin{bmatrix} x_{t-1} - x_{ss} \\ \epsilon_t \\ \chi \end{bmatrix} + \mathcal{H}g^i_{ss}(s_t) \begin{bmatrix} x_{t-1} - x_{ss} \\ \epsilon_t \\ \chi \end{bmatrix}
\]

and for each \( i = 1, \ldots, n_x \):

\[
h^i (x_{t-1}, \epsilon_t, \chi, s_t) - x_{ss} \simeq \mathcal{D}h^i_{ss}(s_t) \begin{bmatrix} x_{t-1} - x_{ss} \\ \epsilon_t \\ \chi \end{bmatrix} + \mathcal{H}h^i_{ss}(s_t) \begin{bmatrix} x_{t-1} - x_{ss} \\ \epsilon_t \\ \chi \end{bmatrix}
\]
where
\[
H^i_{ss} (s_t) = H^i_{ss} (x_{ss}, 0_{n_x}, 0, s_t) = \left[ D_k D_j g^i (x_{ss}, 0_{n_x}, 0, s_t) \right]_{1 \leq j, k \leq n_x + n_x + 1}
\]
is the \((n_x + n_x + 1) \times (n_x + n_x + 1)\) Hessian matrix of \(g^i\) with respect to \((x_{t-1}, \varepsilon_t, \chi)\) evaluated at \((x_{ss}, 0_{n_x}, 0, s_t)\) for all \(s_t\), and
\[
H^i_{ss} (s_t) = H^i_{ss} (x_{ss}, 0_{n_x}, 0, s_t) = \left[ D_k D_j h^i (x_{ss}, 0_{n_x}, 0, s_t) \right]_{1 \leq j, k \leq n_x + n_x + 1}
\]
is the \((n_x + n_x + 1) \times (n_x + n_x + 1)\) Hessian matrix of \(h^i\) with respect to \((x_{t-1}, \varepsilon_t, \chi)\) evaluated at \((x_{ss}, 0_{n_x}, 0, s_t)\) for all \(s_t\).

The objective is now to find the coefficients
\[
\left\{ \left\{ \left\{ \left\{ H^i_{ss} (s) \right\}_{i=1}^{n_x} \right\}_{i=1}^{n_x} \right\}_{s=1}^{n_x} \right\} = \left\{ \left\{ \left\{ D_k D_j g^i (s) \right\}_{i=1}^{n_x} \right\}_{i=1}^{n_x} \right\}_{s=1}^{n_x + n_x + 1} \left\{ \left\{ D_k D_j h^i (s) \right\}_{i=1}^{n_x} \right\}_{i=1}^{n_x + n_x + 1} \right\}_{s=1}^{n_x + n_x + 1}
\]
in the above described expansions.

6 Examples

6.1 Example 1: RBC Model

This section presents a simple exercise to illustrate the theoretical framework at hand. The perfect vehicle for such pedagogical effort is the real business cycle model. There are two reasons. First, the stochastic neoclassical growth model is the foundation of modern macroeconomics. Even the more complicated New Keynesian models, such as those in \cite{?} or \cite{?}, are built around the core of the neoclassical growth model augmented with nominal and real rigidities. Thus, after understanding how to deal with Markov switching in this prototype economy, it will be rather straightforward to extend it to richer environments such as the ones commonly used for policy analysis. Second, the model is so well known, its working so well understood, and its computation so thoroughly explored that the role of time-varying volatility in it will be staggeringly transparent.

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6.2 The RBC Model

Consider a real business cycle model where growth in total factor productivity follows a Markov process with only two regimes. In particular, the TFP process will follow a random walk in logs with drift that takes one of two levels, high and low, so the economy experiences high or low growth. The random walk specification helps simplify the number of variables considered in a stationary equilibrium, and is hence the most parsimonious illustrative example. The specification of two regimes will allow a succinct discussion of the methodology, but, as mentioned above, more regimes can be handled easily within the framework.

To get into the substantive questions as soon as possible, the description of the standard features of the prototype economy will be limited to fixed notation. There is a representative household in the economy, whose preferences over stochastic sequences of consumption, $c_t$, are represented by a utility function:

$$\max_{\mathbb{E}_0} \sum_{t=0}^{\infty} \beta^t \log c_t$$

where $\beta \in (0, 1)$. The resource constraint is

$$c_t + k_t = z_t k_t^{\alpha} + (1 - \delta) k_{t-1}$$

where $k_t$ is capital and the technological change, $z_t$, proceeds according to a random walk in logs with drift where the Markov switching is in the drift, i.e.

$$\log z_t = \mu_t + \log z_{t-1} + \sigma \epsilon_t$$

where the drift takes two values

$$\mu_t = \mu(s_t), \ s_t \in \{1, 2\}$$

and the transition matrix is $P = [p_{i,j}]$ where $p_{i,j} = \Pr(s_t = j | s_{t-1} = i)$.

For this model it is natural to work with the solution to the social planner’s problem. The optimality conditions are standard:

$$\frac{1}{c_t} = \beta \mathbb{E}_t \frac{1}{c_{t+1}} \left( \alpha z_{t+1} k_t^{\alpha-1} + 1 - \delta \right)$$
and
\[ c_t + k_t = z_t k_{t-1}^\alpha + (1 - \delta) k_{t-1}. \]

Due to the unit root the economy is non-stationary. Thus, define \( \omega_t = z_{t-1}^{\frac{1}{\alpha}} \), and let \( \check{c}_t = \frac{c_t}{\omega_t}, \check{k}_{t-1} = \frac{k_{t-1}}{\omega_t}, \check{z}_t = \frac{z_t}{z_{t-1}}. \) Then the re-scaled equilibrium conditions are
\[
\frac{1}{\check{c}_t} = \beta \mathbb{E}_t \frac{\check{z}_t \frac{1}{\alpha}}{\check{c}_{t+1}} \left( \alpha \check{z}_{t+1} \check{k}_{t-1}^{\alpha-1} + 1 - \delta \right),
\]
\[
\check{c}_t + \check{k}_t \check{z}_t^{1-\alpha} = \check{z}_t \check{k}_{t-1}^\alpha + (1 - \delta) \check{k}_{t-1},
\]
and,
\[
\log \check{z}_t = \mu_t + \sigma \varepsilon_t.
\]

Substituting the expression for \( \check{z}_t \), the conditions are then
\[
\frac{1}{\check{c}_t} = \beta \mathbb{E}_t \frac{1}{\check{c}_{t+1}} e^{\frac{\mu_t + \sigma \varepsilon_t}{1-\alpha}} \left( \alpha e^{\mu_{t+1} + \sigma \varepsilon_{t+1}} \check{k}_{t-1}^{\alpha-1} + 1 - \delta \right),
\]
and
\[
\check{c}_t + \check{k}_t e^{\frac{\mu_t + \sigma \varepsilon_t}{1-\alpha}} = e^{\mu_t + \sigma \varepsilon_t} \check{k}_{t-1}^\alpha + (1 - \delta) \check{k}_{t-1}.
\]

Using the notation in Section 2, \( x_{t-1} = \check{k}_{t-1}, y_t = \check{c}_t \), and \( \theta_t = \theta_{t+1} = \mu_t \), so
\[
\mathcal{f} \left( y_{t+1}, y_t, x_t, x_{t-1}, \chi \varepsilon_{t+1}, \varepsilon_t, \theta_{t+1}, \theta_t \right) =
\[
\begin{bmatrix}
\frac{1}{\check{c}_t} - \beta \frac{1}{\check{c}_{t+1}} e^{\frac{\mu_t + \sigma \varepsilon_t}{1-\alpha}} \left( \alpha e^{\mu_{t+1} + \chi \varepsilon_{t+1}} \check{k}_{t-1}^{\alpha-1} + 1 - \delta \right) \\
\check{c}_t + \check{k}_t e^{\frac{\mu_t + \sigma \varepsilon_t}{1-\alpha}} - e^{\mu_t + \sigma \varepsilon_t} \check{k}_{t-1}^\alpha - (1 - \delta) \check{k}_{t-1}
\end{bmatrix}.
\]

Clearly,
\[
\check{c}_t = g \left( \check{k}_{t-1}, \varepsilon_t, \chi, s_t \right) ,
\]
\[
\check{c}_{t+1} = g \left( \check{k}_t, \chi \varepsilon_{t+1}, \chi, s_{t+1} \right) ,
\]
\[
\check{k}_t = h \left( \check{k}_{t-1}, \varepsilon_t, \chi, s_t \right) ,
\]
and the Markov Switching parameter is
\[
\mu_{t+1} = \mu (\chi, s_{t+1}) = \overline{\mu} + \chi \tilde{\mu} (s_{t+1}) .
\]
6.3 Solving the RBC Model

This subsection shows how to solve the model using a first order approximation. The first step is to find the steady state, and the second is to define the matrices in expression (19) that are necessary to solve for the policy functions. Finally, after solving the model, simulations demonstrate the decision rules.

6.3.1 Steady State

In order to calculate steady state, set $\chi = 0$. Therefore, $\tilde{c}_t = \tilde{c}_{t+1} = \tilde{c}_{ss}$, $\tilde{k}_{t-1} = \tilde{k}_t = \tilde{k}_{ss}$, and $\mu_{t+1} = \mu_t = \bar{\mu}$. So the equilibrium conditions in steady state are

$$
\begin{bmatrix}
\frac{1}{\tilde{c}_{ss}} - \frac{\beta}{\bar{\mu}} \frac{1}{\tilde{c}_{ss}} e^{\frac{\bar{\mu}}{\alpha}} \left( \alpha e^{\frac{\bar{\mu}}{\alpha}} \tilde{k}_{ss}^{-1} + 1 - \delta \right)
\tilde{c}_{ss} + \tilde{k}_{ss} e^{\frac{\bar{\mu}}{\alpha}} - e^{\bar{\mu}} \tilde{k}_{ss}^\alpha - (1 - \delta) \tilde{k}_{ss}
\end{bmatrix} = 0_{2 \times 1}
$$

and solve these produces the steady state values

$$
\tilde{k}_{ss} = \left( \frac{1}{\alpha e^{\bar{\mu}}} \left( \frac{1}{\beta e^{\frac{\bar{\mu}}{\alpha}}} - 1 + \delta \right) \right)^{\frac{1}{\alpha-1}}
$$

and

$$
\tilde{c}_{ss} = e^{\bar{\mu}} \tilde{k}_{ss}^\alpha + (1 - \delta) \tilde{k}_{ss} - \tilde{k}_{ss} e^{\frac{\bar{\mu}}{\alpha}}.
$$

6.3.2 The Matrices

The next step is to define the matrices in expression (19), which depend on the derivatives of the function $f$ evaluated at the steady state. Recall in this example that $n_y = 1$, $n_x = 1$, $n_c = 1$, and $n_\theta = 1$. The necessary matrices are

$$
\mathcal{D}_1 f_{ss} (s', s) = \begin{bmatrix} \frac{1}{\tilde{c}_{ss}} \\ 0 \end{bmatrix}, \mathcal{D}_2 f_{ss} (s', s) = \begin{bmatrix} -\frac{1}{\tilde{c}_{ss}} \\ 1 \end{bmatrix}, \mathcal{D}_3 f_{ss} (s', s) = \begin{bmatrix} (1 - \alpha) \alpha \beta e^{\frac{\alpha \bar{\mu}}{\alpha - 1}} \frac{\tilde{k}_{ss}^{\alpha - 2}}{\tilde{c}_{ss}} \\ e^{\frac{\bar{\mu}}{\alpha} - 1} \end{bmatrix}
$$

$$
\mathcal{D}_4 f_{ss} (s', s) = \begin{bmatrix} 0 \\ -e^{\frac{\bar{\mu}}{\beta}} \end{bmatrix}, \mathcal{D}_5 f_{ss} (s', s) = \begin{bmatrix} -\alpha \beta e^{\frac{\alpha \bar{\mu}}{\alpha - 1}} \frac{\tilde{k}_{ss}^{\alpha - 1}}{\tilde{c}_{ss}} \sigma \\ 0 \end{bmatrix},
$$
\[ D_6 f_{ss} (s', s) = \left[ \begin{array}{c} \frac{\sigma}{(1-\alpha)c_{ss}} \\
 \frac{\hat{\epsilon}_s}{\epsilon_{ss}} k_{ss} - \epsilon \mu k_{ss}' \end{array} \right] \]

\[ D_7 f_{ss} (s', s) = \left[ \begin{array}{c} -\alpha \beta \epsilon_{s} - 1 \\
 0 \end{array} \right], \quad D_8 f_{ss} (s', s) = \left[ \begin{array}{c} \frac{1}{c_{ss}} \\
 -\epsilon \mu k_{ss}' + \frac{1}{1-\alpha} \epsilon \frac{\hat{\epsilon}_s}{\epsilon_{ss}} k_{ss}' \end{array} \right] \]

Given these derivatives, constructing the necessary matrices for the solution is straightforward. The first sets of matrices are for the quadratic system, they are

\[
A (1) = \left[ \sum_{s' = 1}^{2} p_{1, s'} D_3 f_{ss} (s', 1) \quad p_{1,1} D_1 f_{ss} (1, 1) \quad p_{1,2} D_1 f_{ss} (2, 1) \right],
\]

\[
A (2) = \left[ \sum_{s' = 1}^{2} p_{2, s'} D_3 f_{ss} (s', 2) \quad p_{2,1} D_1 f_{ss} (1, 2) \quad p_{2,2} D_1 f_{ss} (2, 2) \right],
\]

and

\[
B (1) = - \sum_{s' = 1}^{2} p_{1, s'} \left[ D_4 f_{ss} (s', 1) \quad D_2 f_{ss} (s', 1) \right],
\]

\[
B (2) = - \sum_{s' = 1}^{2} p_{2, s'} \left[ D_4 f_{ss} (s', 2) \quad D_2 f_{ss} (s', 2) \right].
\]

The second set of matrices are used for the derivative with respect to \( \varepsilon_t \), and they are

\[
\Theta_\varepsilon = \sum_{s' = 1}^{2} \left[ \begin{array}{cc} p_{1, s'} D_2 f_{ss} (s', 1) & 0 \\
 0 & p_{2, s'} D_2 f_{ss} (s', 2) \end{array} \right],
\]

\[
\Phi_\varepsilon = \sum_{s' = 1}^{2} \left[ \begin{array}{cc} p_{1, s'} D_1 f_{ss} (s', 1) D_1 g_{ss} (s') + D_3 f_{ss} (s', 1) & 0 \\
 0 & p_{2, s'} D_1 f_{ss} (s', 2) D_1 g_{ss} (s') + D_3 f_{ss} (s', 2) \end{array} \right],
\]

and

\[
\Psi_\varepsilon = - \sum_{s' = 1}^{2} \left[ \begin{array}{cc} p_{1, s'} D_6 f_{ss} (s', 1) \\
 p_{2, s'} D_6 f_{ss} (s', 2) \end{array} \right].
\]

The third set of matrices are used for the derivative with respect to \( \chi \), and they are

\[
\Theta_\chi = \sum_{s' = 1}^{n_s} \left[ \begin{array}{cc} p_{1, s'} D_2 f_{ss} (s', 1) + p_{1,1} D_1 f_{ss} (1, 1) & p_{1,2} D_1 f_{ss} (2, 1) \\
 p_{2,1} D_1 f_{ss} (1, 2) & p_{2, s'} D_2 f_{ss} (s', 2) + p_{2,2} D_1 f_{ss} (2, 2) \end{array} \right],
\]

\[
\Phi_\chi = \sum_{s' = 1}^{2} \left[ \begin{array}{cc} p_{1, s'} D_1 f_{ss} (s', 1) D_1 g_{ss} (s') + p_{1, s'} D_3 f_{ss} (s', 1) & 0 \\
 0 & p_{2, s'} D_1 f_{ss} (s', 2) D_1 g_{ss} (s') + p_{2, s'} D_3 f_{ss} (s', 2) \end{array} \right],
\]
and

$$\Psi_h = - \sum_{s' = 1}^{2} \left[ p_{1,s'} D_{7,8} f_{ss} (s', 1) D \theta_{ss} (s') \right] + \left[ p_{2,s'} D_{7,8} f_{ss} (s', 2) D \theta_{ss} (s') \right].$$

### 6.4 RBC Solution

Now, to describe the solution, first consider the following parameters.

<table>
<thead>
<tr>
<th>$\alpha$</th>
<th>$\beta$</th>
<th>$\delta$</th>
<th>$\sigma$</th>
<th>$\mu$ (1)</th>
<th>$\mu$ (2)</th>
<th>$p_{1,1}$</th>
<th>$p_{2,2}$</th>
</tr>
</thead>
<tbody>
<tr>
<td>0.33</td>
<td>0.99</td>
<td>0.025</td>
<td>0.002</td>
<td>0.03</td>
<td>0.01</td>
<td>0.90</td>
<td>0.90</td>
</tr>
</tbody>
</table>

The transition matrix implies that regimes 1 and 2 occur with equal frequency in the ergodic distribution, so the steady state depends upon $\bar{\mu} = 0.02$. The steady state values of capital and consumption are $k_{ss} = 11.4572$ and $c_{ss} = 1.64771$. Consequently the numerical values of the derivatives are

$$D_1 f_{ss} (s', s) = \begin{bmatrix} 0.3683 \\ 0 \end{bmatrix}, \quad D_2 f_{ss} (s', s) = \begin{bmatrix} -0.3683 \\ 1 \end{bmatrix}, \quad D_3 f_{ss} (s', s) = \begin{bmatrix} 0.0022 \\ 1.0303 \end{bmatrix}$$

$$D_4 f_{ss} (s', s) = \begin{bmatrix} 0 \\ -1.04071 \end{bmatrix}, \quad D_5 f_{ss} (s', s) = \begin{bmatrix} 0 \\ 0 \end{bmatrix}, \quad D_6 f_{ss} (s', s) = \begin{bmatrix} 0.0018 \\ 0.0307 \end{bmatrix}$$

$$D_7 f_{ss} (s', s) = \begin{bmatrix} -0.0383 \\ 0 \end{bmatrix}, \quad \text{and} \quad D_8 f_{ss} (s', s) = \begin{bmatrix} 0.905828 \\ 15.3372 \end{bmatrix}.$$  

Using the Gröbner basis with respect to the ordering $\{D_{1,nz} h_{ss} (1), D_{1,nz} h_{ss} (2), D_{1,nz} g_{ss} (1), D_{1,nz} g_{ss} (2)\}$ produces the following solutions

<table>
<thead>
<tr>
<th>$D_{1,nz} h_{ss} (1)$</th>
<th>$D_{1,nz} g_{ss} (1)$</th>
<th>$D_{1,nz} h_{ss} (2)$</th>
<th>$D_{1,nz} g_{ss} (2)$</th>
</tr>
</thead>
<tbody>
<tr>
<td>1)</td>
<td>1.08526</td>
<td>-0.0774371</td>
<td>1.08526</td>
</tr>
<tr>
<td>2)</td>
<td>0.930745</td>
<td>0.0817605</td>
<td>0.930745</td>
</tr>
<tr>
<td>3)</td>
<td>1.12 - 0.091i</td>
<td>-0.113 + 0.093i</td>
<td>1.12 - 0.091i</td>
</tr>
<tr>
<td>4)</td>
<td>1.12 + 0.091i</td>
<td>-0.113 - 0.093i</td>
<td>1.12 + 0.091i</td>
</tr>
</tbody>
</table>
Now, checking these solutions for MSS, the only stable solution is (2). The full solution is then

\[ s_t = 1: \hat{c}_t = 0.0818\hat{k}_{t-1} + 0.0021\varepsilon_t + 0.0375, \hat{k}_t = 0.9307\hat{k}_{t-1} - 0.0318\varepsilon_t - 0.1852 \]

\[ s_t = 2: \hat{c}_t = 0.0818\hat{k}_{t-1} + 0.0021\varepsilon_t - 0.0375, k_t = 0.9307\hat{k}_{t-1} - 0.0318\varepsilon_t + 0.1852 \]

As an alternative parameterization, consider the same parameters above, but with \( p_{1,1} = 0.5 \). In the ergodic distribution across regimes for this case, regime 1 occurs with probability \( \frac{1}{6} \) and regime 2 occurs with probability \( \frac{5}{6} \). Then the steady state has \( \bar{\mu} = 0.0133, c_{ss} = 1.7967, \) and \( k_{ss} = 14.6326 \), and the first order solution is

\[ s_t = 1: \hat{c}_t = 0.0705\hat{k}_{t-1} + 0.0023\varepsilon_t + 0.0293, \hat{k}_t = 0.9410\hat{k}_{t-1} - 0.0411\varepsilon_t - 0.3526 \]

\[ s_t = 2: \hat{c}_t = 0.0705\hat{k}_{t-1} + 0.0023\varepsilon_t - 0.0058, \hat{k}_t = 0.9410\hat{k}_{t-1} - 0.0411\varepsilon_t + 0.0705 \]

There are two important properties of these first order solutions. First, for both the first case with a symmetric transition matrix and the second case with a non-symmetric transition matrix, the slope coefficients of the solutions are identical across regimes. Second, the additional constant term at the end of the solution is non-zero, which shows the non-certainty equivalence of the first-order solution, and its magnitude depends upon the ergodic probabilities. Since the only regime-switching parameter is the level of growth, the only change in the decision rules is through the constant term, which represent deviations from the steady state due to Markov switching. In the symmetric parameterization, each regime occurs with equal probability in the ergodic distribution, so the steady state is exactly between each regime, and hence the deviations are equally above and below. In the non-symmetric transition matrix parameterization, since regime 2 occurs with a higher probability in the ergodic distribution, the additional constants are much smaller for regime 2, demonstrating that the steady state is closer to regime 2.

Figure ?? shows the policy functions for each regime when the transition matrix is symmetric if \( \varepsilon_t = 0 \), alongside the fixed regime case, which is no Markov switching but with TFP growth always at \( \bar{\mu} \). The plot shows how the policy functions with Markov switching have identical slopes to those without switching, but the constant term associated with Markov switching scales the functions up and down. In the case with a symmetric transition matrix and hence
equal ergodic probabilities, the fixed regime case lies exactly between the two lines when there is switching.

Figure ?? shows the policy functions for the non-symmetric transition matrix case. Again, this figure shows that the slopes are the same, but the Markov switching rules are scaled up and down by a constant. Since regime 2 occurs with higher probability in the ergodic distribution, the fixed regime policy function is very close to that for regime 2, while regime 1 is farther away.

6.5 RBC Simulations

To illustrate how Markov switching can play a role in growth dynamics, especially through the non-certainty equivalence of the first-order approximation, Figures ?? and ?? show simulation results of the models discussed above and their ergodic distributions. For both the symmetric and non-symmetric transition matrices, there are 1000 simulations of the economy for a length of 10000 periods, excluding the first 1000 to eliminate the effects of initial conditions.

Figure ?? shows the simulated distributions of output and consumption growth for the symmetric transition matrix economy. Recall that in this specification, both regimes are equally highly persistent, so in the ergodic distribution, both occur with equal probability. While the fixed regime case has a single-peaked distribution, thereby exhibiting growth at close to a constant rate, the switching case has a twin-peaked distribution for both variables, one peak associated with each regime. The parameterization for the fixed regime case suggests that its growth rate peak should be halfway between the growth rates of the two regimes, but simulations show that growth is higher on average in the switching case than the fixed regime case. This result follows from the non-certainty equivalence of the solution; when there is switching between high and low growth regimes, agents understand that they will experience both regimes, and, on average, this decision leads to higher consumption and output growth than if there was only a single regime.

Figure ?? shows the simulated distributions of output and consumption growth for the case of the non-symmetric transition matrix. In this case, regime 2 occurs much more often and is more persistent than regime 1, so the ergodic distribution has higher probability on regime
2. Again, the fixed regime case exhibits almost constant growth: the distribution is single-peaked. The Markov switching case, on the other hand, is no longer twin-peaked. In this case, there is one dominant peak of the distribution, which is associated with regime 2, but there are also several other smaller peaks to the distribution that correspond to different histories of the regimes. For example, the large peak is a result of regime 2 occurring approximately 5/6 of the time, but there will be long periods where only regime 2 occurs, and the small left-most peak is associated with these stretches. The other smaller peaks correspond to various lengths of regime 1 occurring, which happens with lower probability. As in the symmetric case, the ergodic mean of growth in the fixed case is lower than when there is switching, again this is because of the lack of certainty equivalence in the two regimes.

6.6 Example 2: NK Model

This section presents a second example: a simple New Keynesian model to highlight the issue of determinacy and mean square stability.

6.7 The NK Model

The model is a model with quadratic price adjustment costs where the monetary authority follows a Taylor Rule that changes according to a Markov Process. The reaction coefficient of monetary policy switches with the regime, which ?, ?, and ?, among others, have argued captures the changing stance of policy in the United States.

A representative consumer maximizes expected lifetime utility over consumption \( C_t \) and hours worked \( H_t \)

\[
\mathbb{E}_0 \sum_{t=0}^{\infty} \beta^t (\log C_t - H_t)
\]

subject to the budget constraint

\[
C_t + \frac{B_t}{P_t} = W_t H_t + R_{t-1} \frac{B_{t-1}}{P_{t-1}} + T_t + D_t
\]

where \( B_t \) is next period’s nominal bonds, \( W_t \) is the real wage, \( R_{t-1} \) is the nominal return on bonds, \( T_t \) is lump-sum transfers, and \( D_t \) is profits from firms.
A competitive final goods producer combines a continuum of intermediate goods $Y_{j,t}$ into a final good $Y_t$ by a CES aggregator

$$Y_t = \left( \int_0^1 Y_{j,t}^{\frac{n-1}{n}} dj \right)^{\frac{n}{n-1}}$$

Intermediate goods firms take the wage and their demand function

$$Y_{j,t} = \left( \frac{P_{j,t}}{P_t} \right)^{-\eta} Y_t$$

as given and set their price $P_{j,t}$ demand hours $H_{j,t}$ to produce according to

$$Y_{j,t} = A_t H_{j,t}$$

where total factor productivity follows

$$\log A_t = \mu_t + \log A_{t-1}$$

where, similar to the RBC model in Section 6.1, the drift can take two values

$$\mu_t = \mu(s_t), \ s_t \in \{1, 2\}.$$ 

These firms face quadratic price adjustment costs according to

$$AC_{j,t} = \frac{\kappa}{2} \left( \frac{P_{j,t}}{P_{j,t-1}} - 1 \right)^2.$$ 

The monetary authority sets prices by a Taylor rule where the coefficient varies over time

$$\frac{R_t}{R_{ss}} = \left( \frac{R_{t-1}}{R_{ss}} \right)^{\rho} \Pi_t^{(1-\rho)\psi_t} \exp(\sigma \varepsilon_t)$$

In a symmetric equilibrium $P_{j,t} = P_t$, $Y_{j,t} = Y_t$, and $H_{j,t} = H_t$ for all $j$, and market clearing implies

$$Y_t = C_t + \frac{\kappa}{2} (\Pi_t - 1)^2 Y_t.$$ 

Using the notation in Section 2, $y_t = [\Pi_t, Y_t']$, $x_{t-1} = R_{t-1}$, $\theta_{1t} = \mu_t$, and $\theta_{2t} = \psi_t$. Then the stationary equilibrium is expressed as

$$f(y_{t+1}, y_t, x_t, x_{t-1}, \chi \varepsilon_{t+1}, \varepsilon_t, \theta_{t+1}, \theta_t) =$$

$$\begin{bmatrix}
1 - \beta \frac{\left(1 - \frac{\kappa}{2} (\Pi_t - 1)^2\right) Y_t}{\left(1 - \frac{\kappa}{2} (\Pi_{t+1} - 1)^2\right) Y_{t+1} \exp(\mu_{t+1})} \frac{R_t}{R_{t+1}} \\
(1 - \eta) + \eta \left(1 - \frac{\kappa}{2} (\Pi_t - 1)^2\right) \tilde{Y}_t + \beta \frac{\left(1 - \frac{\kappa}{2} (\Pi_t - 1)^2\right)}{\left(1 - \frac{\kappa}{2} (\Pi_{t+1} - 1)^2\right)} (\Pi_{t+1} - 1) \Pi_{t+1} - \kappa (\Pi_t - 1) \Pi_t \\
\left(\frac{R_{t-1}}{R_{ss}}\right)^\rho \Pi_t^{(1-\rho)\psi_t} \exp(\sigma \varepsilon_t) - \frac{R_t}{R_{ss}}
\end{bmatrix}$$

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6.8 Solving the NK Model

Similar to the RBC example, this subsection defines the steady state and matrices for the first order approximation for the NK example.

6.8.1 Steady State

In order to calculate steady state, set \( \chi = 0 \). Therefore, \( \Pi_t = \Pi_{t+1} = \Pi_{ss} \), \( \bar{Y}_t = \bar{Y}_{t+1} = \bar{Y}_{ss} \), \( R_t = R_{t-1} = R_{ss} \), and \( \mu_{t+1} = \mu_t = \bar{\mu} \). So the equilibrium conditions in steady state are

\[
\begin{bmatrix}
1 - \beta \frac{(1 - \frac{\kappa}{2} (\Pi_{ss} - 1)^2)}{(1 - \frac{\kappa}{2} (\Pi_{ss} - 1)^2)} \bar{Y}_{ss} - \frac{1}{\Pi_{ss}} R_{ss} \\
\eta (1 - \frac{\kappa}{2} (\Pi_{ss} - 1)^2) \bar{Y}_{ss} + \beta \kappa \frac{(1 - \frac{\kappa}{2} (\Pi_{ss} - 1)^2)}{(1 - \frac{\kappa}{2} (\Pi_{ss} - 1)^2)} (\Pi_{ss} - 1) \Pi_{ss} - \kappa (\Pi_{ss} - 1) \Pi_{ss}
\end{bmatrix}
= 0_{3 \times 1}
\]

Using the assumption \( \Pi_{ss} = 1 \), and solving these produces the steady state values

\[
R_{ss} = \frac{\exp (\bar{\mu})}{\beta},
\]

and

\[
\bar{Y}_{ss} = \frac{\eta - 1}{\eta}.
\]

Note that \( \bar{\mu} \) affects the steady state, but \( \psi (s) \) does not, demonstrating the partition of the switching vector \( \theta_t = [\theta_{1t}, \theta_{2t}] \).

6.8.2 The Matrices

The next step is to define the matrices in expression (19), which depend on the derivatives of the function \( f \) evaluated at the steady state. Recall in this example that \( n_y = 2, n_x = 1, n_c = 1, \) and \( n_{\theta} = 2 \). The necessary matrices are

\[
\mathcal{D}_{1,2} f_{ss} (s', s) = \begin{bmatrix}
\frac{\eta}{\eta - 1} & 1 \\
0 & \beta \kappa
\end{bmatrix}, \quad \mathcal{D}_{3,4} f_{ss} (s', s) = \begin{bmatrix}
\frac{\eta}{1 - \eta} & 0 \\
\eta & -\kappa
\end{bmatrix}, \quad \mathcal{D}_{5} f_{ss} (s', s) = \begin{bmatrix}
-\beta e^{-\bar{\mu}} \\
0
\end{bmatrix},
\]

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6.9 NK Solution

The calibration used is as follows

<table>
<thead>
<tr>
<th>β</th>
<th>κ</th>
<th>η</th>
<th>ρ</th>
<th>σ</th>
<th>p_{1,1}</th>
<th>p_{2,2}</th>
<th>\bar{\mu}</th>
<th>\mu(1)</th>
<th>\mu(2)</th>
<th>\psi(1)</th>
<th>\psi(2)</th>
</tr>
</thead>
<tbody>
<tr>
<td>0.99</td>
<td>161</td>
<td>10</td>
<td>0.8</td>
<td>0.0025</td>
<td>0.90</td>
<td>0.90</td>
<td>0.02</td>
<td>0.03</td>
<td>0.01</td>
<td>3.1</td>
<td>0.9</td>
</tr>
</tbody>
</table>

The transition matrix implies that regimes 1 and 2 occur with equal frequency in the ergodic distribution, so the steady state depends upon \( \bar{\mu} = 0.02 \). The steady state values of output and consumption are \( R_{ss} = 1.03051 \) and \( \bar{Y}_{ss} = 0.90 \). Consequently the numerical values of the derivatives are

\[
\mathcal{D}_{1,2}f_{ss}(s', s) = \begin{bmatrix} 1.11111 & 1 \\ 0 & 159.39 \\ 0 & 0 \end{bmatrix}, \quad \mathcal{D}_{3,4}f_{ss}(s', s) = \begin{bmatrix} -1.11111 & 0 \\ 10 & -161. \\ 0 & 0.2\psi(s) \end{bmatrix},
\]

\[
\mathcal{D}_5f_{ss}(s', s) = \begin{bmatrix} -0.970397 \\ 0 \\ -0.970397 \end{bmatrix}, \quad \mathcal{D}_6f_{ss}(s', s) = \begin{bmatrix} 0 \\ 0 \\ 0.776317 \end{bmatrix}, \quad \mathcal{D}_7f_{ss}(s', s) = \begin{bmatrix} 0 \\ 0 \\ 0 \end{bmatrix},
\]

\[
\mathcal{D}_8f_{ss}(s', s) = \begin{bmatrix} 0 \\ 0 \\ 0.0025 \end{bmatrix}, \quad \mathcal{D}_{9,10}f_{ss}(s', s) = \begin{bmatrix} 1 \\ 0 \\ 0 \end{bmatrix}, \quad \mathcal{D}_{11,12}f_{ss}(s', s) = \begin{bmatrix} 0 \\ 0 \\ 0 \end{bmatrix}
\]

Using (19) and the described calibration produces a quadratic system to be solved to find \( \{\mathcal{D}_{1,n_x}g_{ss}(s), \mathcal{D}_{1,n_x}h_{ss}(s)\}_{s=1}^{n_x} \). Using the Gröbner basis with respect to the ordering

\[
\{\mathcal{D}_{1,n_x}h_{ss}(1), \mathcal{D}_{1,n_x}h_{ss}(2), \mathcal{D}_{1,n_x}g_{ss}(1), \mathcal{D}_{1,n_x}g_{ss}(2)\}
\]
the solutions are

<table>
<thead>
<tr>
<th></th>
<th>$D_{1,n_x}h_{ss} (1)$</th>
<th>$D_{1,n_x}g_{ss} (1)'$</th>
<th>$D_{1,n_x}h_{ss} (2)$</th>
<th>$D_{1,n_x}g_{ss} (2)'$</th>
</tr>
</thead>
<tbody>
<tr>
<td>1)</td>
<td>0.596</td>
<td>-1.892</td>
<td>0.700</td>
<td>-2.892</td>
</tr>
<tr>
<td>2)</td>
<td>0.777</td>
<td>-3.575</td>
<td>1.308</td>
<td>-7.266</td>
</tr>
<tr>
<td>3)</td>
<td>0.799</td>
<td>-1.757</td>
<td>1.055</td>
<td>1.332</td>
</tr>
<tr>
<td>4)</td>
<td>1.096-0.438i</td>
<td>-0.791+4.136i</td>
<td>0.463+0.685i</td>
<td>1.337+0.0569i</td>
</tr>
<tr>
<td>5)</td>
<td>1.096+0.438i</td>
<td>-0.791-4.136i</td>
<td>0.463+0.685i</td>
<td>1.337-0.0569i</td>
</tr>
<tr>
<td>6)</td>
<td>1.098-0.208i</td>
<td>-0.963+1.862i</td>
<td>0.467-0.325i</td>
<td>1.026-0.019i</td>
</tr>
<tr>
<td>7)</td>
<td>1.098+0.208i</td>
<td>-0.963-1.862i</td>
<td>0.467+0.325i</td>
<td>1.026+0.019i</td>
</tr>
<tr>
<td>8)</td>
<td>1.240-0.250i</td>
<td>0.756+2.978i</td>
<td>0.688-0.392i</td>
<td>0.752+0.005i</td>
</tr>
<tr>
<td>9)</td>
<td>1.240+0.250i</td>
<td>0.756-2.978i</td>
<td>0.688+0.392i</td>
<td>0.752-0.005i</td>
</tr>
</tbody>
</table>

Checking these solutions for MSS, the first solution is the only stable one. Constructing the full solution produces:

\[
\begin{align*}
\mathbf{s}_t &= 1: \\
&\begin{bmatrix}
\hat{R}_t \\
\hat{Y}_t \\
\hat{\Pi}_t
\end{bmatrix} = \\
&\begin{bmatrix}
0.5965 \\
-1.8919 \\
-0.3184
\end{bmatrix} + \\
&\begin{bmatrix}
0.0019 \\
-0.0062 \\
-0.0010
\end{bmatrix} \varepsilon_t + \\
&\begin{bmatrix}
-0.0014 \\
0.0250 \\
-0.0022
\end{bmatrix}
\end{align*}
\]

\[
\begin{align*}
\mathbf{s}_t &= 2: \\
&\begin{bmatrix}
\hat{R}_t \\
\hat{Y}_t \\
\hat{\Pi}_t
\end{bmatrix} = \\
&\begin{bmatrix}
0.7004 \\
-2.8919 \\
-0.5366
\end{bmatrix} + \\
&\begin{bmatrix}
0.0022 \\
-0.0095 \\
-0.0018
\end{bmatrix} \varepsilon_t + \\
&\begin{bmatrix}
-0.0043 \\
-0.0724 \\
-0.0230
\end{bmatrix}
\end{align*}
\]

where \(\var_t = \text{var} - \text{var}_{ss}\) where \(\text{var}_{ss}\) is the steady state of \(\text{var}\) for \(\text{var} \in \{R_t, Y_t, \Pi_t\}\).

As an alternative, suppose now that \(\psi (2) = 0.7\). There are still nine total solutions, but now there are two stable solutions:

<table>
<thead>
<tr>
<th></th>
<th>$D_{1,n_x}h_{ss} (1)$</th>
<th>$D_{1,n_x}g_{ss} (1)'$</th>
<th>$D_{1,n_x}h_{ss} (2)$</th>
<th>$D_{1,n_x}g_{ss} (2)'$</th>
</tr>
</thead>
<tbody>
<tr>
<td>1)</td>
<td>0.592109</td>
<td>-1.90874</td>
<td>0.713454</td>
<td>-3.14857</td>
</tr>
<tr>
<td>2)</td>
<td>0.858767</td>
<td>-1.70451</td>
<td>1.01631</td>
<td>2.13138</td>
</tr>
</tbody>
</table>

which shows that this parameterization does not produce a unique MSS solution.
These two parameterizations demonstrate how MSS as a stability concept determines existence and uniqueness of the solution. In the parameterization with $\psi(2) = 0.9$, solving the system produces 9 total solutions, and only satisfies mean square stability. In the case of $\psi(2) = 0.7$, having two MSS solutions implies non-uniqueness of a stable solution. If, on the other hand, there were no MSS solutions, then a stable solution does not exist.

6.10 NK Model with Habits

Now consider a slight variant of the previously discussed New Keynesian model, but with households that have habit formation. In this case, they maximize

$$\mathbb{E}_0 \sum_{t=0}^{\infty} \beta^t (\log (C_t - \varphi C_{t-1}) - H_t)$$

where $\varphi$ denotes habit persistence. For simplicity, assume TFP is constant

$$A_t = 1$$

and that there is no interest rate smoothing, so $\rho = 0$:

$$\frac{R_t}{R_{ss}} = \Pi^e_t \exp (\sigma \varepsilon_{r,t}).$$

With habits, consumption appears dated as $C_{t-1}$, $C_t$, and $C_{t+1}$ in the equilibrium conditions, which are

$$\lambda_t = \frac{1}{C_t - \varphi C_{t-1}} - \beta \mathbb{E}_t \frac{\varphi}{C_{t+1} - \varphi C_t}$$

$$\lambda_t = \beta \mathbb{E}_t \lambda_{t+1} \frac{R_t}{\Pi_{t+1}}$$

$$\kappa \lambda_t (\Pi_t - 1) \Pi_t = (1 - \eta) \lambda_t + \eta + \beta \kappa \mathbb{E}_t \lambda_{t+1} (\Pi_{t+1} - 1) \Pi_{t+1} \frac{Y_{t+1}}{Y_t}$$

$$\frac{R_t}{R_{ss}} = \Pi^e_t \exp (\sigma \varepsilon_{r,t})$$

$$Y_t = C_t + \frac{\kappa}{2} (\Pi_t - 1)^2 Y_t$$
Consequently, define the auxiliary variable $\tilde{C}_t = C_t$, so $\tilde{C}_{t+1} = C_{t+1}$. Substituting out $R_t$ and $Y_t$ to simplify the equilibrium conditions, now gives $y_t = \left[ \Pi_t, \tilde{C}_t, \lambda_t \right]', y_{t+1} = \left[ \Pi_{t+1}, \tilde{C}_{t+1}, \lambda_{t+1} \right]$, $x_t = [C_t]$, $x_{t-1} = [C_{t-1}]'$, and $\theta_{2t} = \psi_t$. Then the equilibrium is expressed as

$$f (y_{t+1}, y_t, x_t, x_{t-1}, \chi \varepsilon_{t+1}, \varepsilon_t, \theta_{t+1}, \theta_t) =$$

$$= \frac{1}{\tilde{C}_t - \varphi C_{t-1}} - \beta E_t \frac{\varphi}{\tilde{C}_{t+1} - \varphi C_t} - \lambda_t$$

$$- \beta E_t \lambda_{t+1} \frac{R_{ss} \Pi_t \varepsilon_t \exp(\sigma \varepsilon_{t+1})}{\Pi_{t+1}} - \lambda_t$$

$$(1 - \eta) \lambda_t + \eta + \beta \kappa E_t \lambda_{t+1} (\Pi_{t+1} - 1) \frac{C_{t+1}}{C_t} \frac{1 - \frac{\varphi}{\varphi}}{1 - \frac{\varphi}{\varphi}} - \kappa \lambda_t (\Pi_t - 1) \Pi_t$$

$$\tilde{C}_t - C_t$$

Assuming $\Pi_{ss} = 1$, the steady state satisfies

$$\lambda_{ss} = \frac{\eta}{\eta - 1}$$

$$C_{ss} = \frac{1 - \beta \varphi \eta - 1}{1 - \varphi \eta}$$

$$\tilde{C}_{ss} = C_{ss}$$

Taking derivatives with respect to the vector $\left[ \Pi_{t+1}, \tilde{C}_{t+1}, \lambda_{t+1}, \Pi_t, \tilde{C}_t, \lambda_t, C_t, C_{t-1}, \varepsilon_{t+1}, \varepsilon_t, \psi_{t+1}, \psi_t \right]$ and evaluating at the steady state produces

$$\mathcal{D} f_{ss} (s', s) =$$

$$= \begin{bmatrix}
0 & \frac{\varphi^2}{c_{ss}^2 (1 - \varphi)} & 0 & 0 & 0 & -1 - \frac{1 + \beta \varphi^2}{c_{ss}^2 (1 - \varphi)^2} & \frac{\varphi}{c_{ss}^2 (1 - \varphi)^2} & 0 & 0 & 0 & 0 \\
-\lambda_{ss} & 0 & 1 & \lambda_{ss} \psi' (s) & 0 & -1 & 0 & 0 & 0 & \lambda_{ss} \sigma & 0 & 0 \\
-\beta \kappa \lambda_{ss} & 0 & 0 & -\kappa \lambda_{ss} & 0 & 1 - \eta & 0 & 0 & 0 & 0 & 0 & 0 \\
0 & 0 & 0 & 0 & 0 & 1 & 0 & -1 & 0 & 0 & 0 & 0
\end{bmatrix}$$

The calibration used is as follows

$$\begin{array}{cccccccc}
\beta & \kappa & \eta & \varphi & \sigma & p_{1,1} & p_{2,2} \\
0.99 & 161 & 10 & 0.7 & 0.0025 & 0.90 & 0.90
\end{array}$$
and so the numerical value of the derivative matrix is

\[
D_{f_{ss}}(s', s) = \begin{bmatrix}
0 & 9.0776 & 0 & 0 & 0 & -1 & -19.453 & 9.1693 & 0 & 0 & 0 & 0 \\
-1.1111 & 0 & 1 & 1.111\psi(s) & 0 & -1 & 0 & 0 & 0 & 0.0028 & 0 & 0 \\
177.1 & 0 & 0 & -178.89 & 0 & -9.0 & 0 & 0 & 0 & 0 & 0 & 0 \\
0 & 0 & 0 & 0 & 1 & 0 & -1 & 0 & 0 & 0 & 0 & 0
\end{bmatrix}
\]

Given that there are \( n_y = 3 \) nonpredetermined variables, \( n'_x = 1 \) endogenous predetermined variable, and \( n_s = 2 \) regimes, the there are

\[
\begin{bmatrix}
\begin{bmatrix} n_{\text{endo}} \end{bmatrix}^{n_y} \\
\begin{bmatrix} n'_x \end{bmatrix}
\end{bmatrix} = \left( \frac{4!}{1!3!} \right)^2 = 16
\]

total solutions.

If \( \psi(1) = 1.1 \) and \( \psi(2) = 0.7 \), there are two MSS solutions. They are, in part

\[
\begin{array}{c|cc}
 & D_{1,n_x,h_{ss}}(1) & D_{1,n_x,h_{ss}}(2) \\
1) & 0.7 & 0.7 \\
2) & 0.76566 & 0.9810 \\
\end{array}
\]

If \( \psi(1) = 3.1 \) and \( \psi(2) = 0.7 \), there is one MSS solution, and it is, for \( s_t = 1 \)

\[
\begin{align*}
\hat{y}_t &= \begin{bmatrix} 0 \\ 0.7 \\ 0 \end{bmatrix} + \begin{bmatrix} -0.00012 \\ -0.00018 \\ 0.00239 \end{bmatrix} \varepsilon_t + \begin{bmatrix} 0 \\ 0 \\ 0 \end{bmatrix} \\
\hat{x}_t &= 0.7\hat{x}_{t-1} - 0.00018\varepsilon_t + 0
\end{align*}
\]

and for \( s_t = 2 \):

\[
\begin{align*}
\hat{y}_t &= \begin{bmatrix} 0 \\ 0.7 \\ 0 \end{bmatrix} + \begin{bmatrix} -0.00013 \\ -0.00020 \\ 0.00264 \end{bmatrix} \varepsilon_t + \begin{bmatrix} 0 \\ 0 \\ 0 \end{bmatrix} \\
\hat{x}_t &= 0.7\hat{x}_{t-1} - 0.00020\varepsilon_t + 0
\end{align*}
\]
7 Conclusion

This paper developed a perturbation method for constructing approximations to the solutions to Markov switching DSGE models. The framework allows introducing switching from first principles, including when switching affects the steady state of the economy. While not pursued here, second- or higher-order approximations are straightforward, and follow the single-regime case studied by ?. Using Gröbner bases to solve the system and mean square stability to check for existence and uniqueness of stable solutions, the method handles a wide variety of models, and shows that switching in parameters that affect the steady state implies that first order approximations are not certainty equivalent.

8 Appendix: Second Order Derivatives

\[ \mathcal{H}_{1,n_x;1,n_z} G^i_{ss} (s_t) = \sum_{s'=1}^{n_s} p_{s,t,s'} \times \]

\[
\begin{pmatrix}
[D_{1,n_x} g_{ss} (s') D_{1,n_x} h_{ss} (s_t)]^T T + D_{1,n_x} g_{ss} (s_t) \times \\
D_{1,n_x} h_{ss} (s_t) \times \\
+ D_{1,n_x} g_{ss} (s_t) \times \\
+ D_{1,n_x} h_{ss} (s_t) \times \\
+ \mathcal{H}_{n_y+1,2n_y+n_x;2n_y+1,2n_y+n_x} f^i_{ss} (s', s_t) \mathcal{H}_{1,n_x;1,n_z} g_{ss} (s') \times \\
+ \mathcal{H}_{n_y+1,2n_y+n_x;2n_y+1,2n_y+n_x} f^i_{ss} (s', s_t) \mathcal{H}_{1,n_x;1,n_z} h_{ss} (s_t) \times \\
+ \mathcal{H}_{2n_y+1,2n_y+n_x;2n_y+1,2n_y+n_x} f^i_{ss} (s', s_t) \mathcal{H}_{1,n_x;1,n_z} g_{ss} (s') \times \\
+ \mathcal{H}_{2n_y+1,2n_y+n_x;2n_y+1,2n_y+n_x} f^i_{ss} (s', s_t) \mathcal{H}_{1,n_x;1,n_z} h_{ss} (s_t) \times \\
\end{pmatrix}
\]
\[
\mathcal{H}_{1,n_x:n_x+1,n_x+n_c} \mathcal{G}_{ss}^i (s_t) = \sum_{s'=1}^{n_x} p_{s', s_t} \times \\
\begin{pmatrix}
\mathcal{D}_{1,n_x} g_{ss} (s_t) \\
\mathcal{D}_{1,n_x} h_{ss} (s_t)
\end{pmatrix}^T
\begin{bmatrix}
\mathcal{H}_{1,n_y:1,n_y} f_{ss}^i (s', s_t) \mathcal{D}_{1,n_x} g_{ss} (s') + 2\mathcal{H}_{1,n_y:2n_y+1,2n_y+n_x} f_{ss}^i (s', s_t) \\
+ \mathcal{H}_{1,n_y:n_y+1,2n_y+n_x} f_{ss}^i (s', s_t) \mathcal{D}_{n_x+1,n_x+n_c} h_{ss} (s_t)
\end{bmatrix}
\begin{bmatrix}
\mathcal{D}_{1,n_x} h_{ss} (s_t) \\
\mathcal{D}_{1,n_x} g_{ss} (s_t)
\end{bmatrix}^T
\begin{bmatrix}
\mathcal{H}_{n_y+1,2n_y+1,n_y} f_{ss}^i (s', s_t) \mathcal{D}_{1,n_x} g_{ss} (s') + \mathcal{H}_{n_y+1,2n_y+2n_y+1,2n_y+n_x} f_{ss}^i (s', s_t) \\
+ \mathcal{H}_{n_y+1,2n_y+2n_y+1,2n_y+n_x} f_{ss}^i (s', s_t) \mathcal{D}_{n_x+1,n_x+n_c} g_{ss} (s_t)
\end{bmatrix}
\begin{bmatrix}
\mathcal{D}_{n_x+1,n_x+n_c} h_{ss} (s_t) \\
\mathcal{D}_{n_x+1,n_x+n_c} h_{ss} (s_t)
\end{bmatrix}^T
\begin{bmatrix}
\mathcal{H}_{2n_y+1,2n_y+2n_y+2n_y+1,2n_y+n_x} f_{ss}^i (s', s_t) \\
+ \mathcal{D}_{1,n_y} f_{ss}^i (s', s_t) \mathcal{H}_{1,n_x:1,n_x} g_{ss} (s') \\
+ \mathcal{D}_{1,n_y} f_{ss}^i (s', s_t) \mathcal{H}_{1,n_x:1,n_x} g_{ss} (s') \mathcal{D}_{n_x+1,n_x+n_c} h_{ss} (s_t)
\end{bmatrix}
\begin{bmatrix}
\mathcal{D}_{1,n_y} f_{ss}^i (s', s_t) \mathcal{D}_{1,n_x} g_{ss} (s') + \mathcal{D}_{2n_y+1,2n_y+2n_y+1,2n_y+n_x} f_{ss}^i (s', s_t) \\
+ \mathcal{H}_{2n_y+1,2n_y+2n_y+1,2n_y+n_x} f_{ss}^i (s', s_t) \mathcal{D}_{n_x+1,n_x+n_c} g_{ss} (s_t)
\end{bmatrix}
\begin{bmatrix}
\mathcal{D}_{1,n_x} h_{ss} (s_t) \\
\mathcal{D}_{n_x+1,n_x+n_c} h_{ss} (s_t)
\end{bmatrix}^T
\begin{bmatrix}
\mathcal{D}_{1,n_x} g_{ss} (s_t) + \mathcal{D}_{2n_y+1,2n_y+2n_y+1,2n_y+n_x} f_{ss}^i (s', s_t) \\
+ \mathcal{H}_{2n_y+1,2n_y+2n_y+1,2n_y+n_x} f_{ss}^i (s', s_t) \mathcal{D}_{n_x+1,n_x+n_c} g_{ss} (s_t)
\end{bmatrix}
\begin{bmatrix}
\mathcal{D}_{n_x+1,n_x+n_c} h_{ss} (s_t) \\
\mathcal{D}_{n_x+1,n_x+n_c} h_{ss} (s_t)
\end{bmatrix}^T
\begin{bmatrix}
\mathcal{D}_{n_x+1,n_x+n_c} h_{ss} (s_t) \\
\mathcal{D}_{n_x+1,n_x+n_c} h_{ss} (s_t)
\end{bmatrix}^T
\end{bmatrix}
\]
\[ H_{1,n:x+n_x+1}f_{ss}^i (s_t) = \sum_{s'=1}^{n_s} p_{s_t,s'} \times \]

\[
\begin{bmatrix}
H_{1,n:y;1,n:y}f_{ss}^i (s', s_t) \\
H_{1,n:y;1,n:y}g_{ss}^i (s', s_t)
\end{bmatrix}^T
\begin{bmatrix}
D_{n_x+n_x+1}g_{ss} (s') \\
D_{n_x+n_x+1}h_{ss} (s_t)
\end{bmatrix}
+ D_{1,n:y}g_{ss} (s_t) T
\]

\[
[ D_{1,n:x}g_{ss} (s') D_{1,n:x}h_{ss} (s_t) ]^T
\]

\[
\begin{bmatrix}
D_{n_x+n_x+1}g_{ss} (s') \\
D_{n_x+n_x+1}h_{ss} (s_t)
\end{bmatrix}
\]

\[
H_{2n_y+1,2n_y;1,n:y}f_{ss}^i (s', s_t) [ D_{n_x+n_x+1}g_{ss} (s') + D_{1,n:x}g_{ss} (s') D_{n_x+n_x+1}h_{ss} (s_t) ]
\]

\[
+ H_{n:y+1,2n_y;1,n:y}f_{ss}^i (s', s_t) D_{n_x+n_x+1}g_{ss} (s')
\]

\[
+ H_{n:y+1,2n_y;1,n:y}g_{ss}^i (s', s_t) D_{n_x+n_x+1}h_{ss} (s_t)
\]

\[
H_{n:y+1,2n_y;1,n:y}f_{ss}^i (s', s_t)
\]

\[
+ H_{n:y+1,2n_y;1,n:y}g_{ss}^i (s', s_t) D_{n_x+n_x+1}g_{ss} (s')
\]

\[
+ H_{n:y+1,2n_y;1,n:y}g_{ss}^i (s', s_t) D_{n_x+n_x+1}h_{ss} (s_t)
\]

\[
+ H_{n:y+1,2n_y;1,n:y}f_{ss}^i (s', s_t) D_{n_x+n_x+1}g_{ss} (s')
\]

\[
+ H_{n:y+1,2n_y;1,n:y}g_{ss}^i (s', s_t) D_{n_x+n_x+1}h_{ss} (s_t)
\]

\[
+ H_{n:y+1,2n_y;1,n:y}f_{ss}^i (s', s_t) D_{n_x+n_x+1}g_{ss} (s')
\]

\[
+ H_{n:y+1,2n_y;1,n:y}g_{ss}^i (s', s_t) D_{n_x+n_x+1}h_{ss} (s_t)
\]

\[
+ H_{n:y+1,2n_y;1,n:y}f_{ss}^i (s', s_t) D_{n_x+n_x+1}g_{ss} (s')
\]

\[
+ H_{n:y+1,2n_y;1,n:y}g_{ss}^i (s', s_t) D_{n_x+n_x+1}h_{ss} (s_t)
\]

\[
+ [ D_{1,n:y}f_{ss}^i (s', s_t) D_{1,n:x}g_{ss} (s') + D_{2n_y+1,2n_y+n_x}f_{ss}^i (s', s_t) ] H_{1,n:x;1,n:x+2}g_{ss} (s_t)
\]

\[
+ D_{n:y+1,2n_y}f_{ss}^i (s', s_t) H_{1,n:x;1,n:x+1}g_{ss} (s_t)
\]

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\[ \mathcal{H}_{n_x+1,n_x+n_z;n_x+1,n_x+n_z} G_{ss}^i (s_t) = \sum_{s'=1}^{n_s} p_{s_t,s'} \times \]

\[
\begin{pmatrix}
[D_{n_x} g_{ss} (s') D_{n_x+1,n_x+n_z} h_{ss} (s_t)] & [H_{1,n_y;1,n_y} f_{ss}^i (s', s_t) D_{1,n_x} g_{ss} (s') + 2H_{1,n_y;2n_y+1,2n_y+n_x} f_{ss}^i (s', s_t)] & [D_{n_x+1,n_x+n_z} h_{ss} (s_t)] \\
+D_{n_x+1,n_x+n_z} g_{ss} (s_t)^T & \left[ H_{2n_y+1,2n_y+n_x;2n_y+1,2n_y+n_x} f_{ss}^i (s', s_t) \right] & \left[ +D_{1,n_y} f_{ss}^i (s', s_t) H_{1,n_x;1,n_x} g_{ss} (s') \right] & [D_{n_x+1,n_x+n_z} h_{ss} (s_t)] \\
+D_{n_x+1,n_x+n_z} h_{ss} (s_t)^T & \left[ H_{n_y+1,2n_y+n_x;1,2n_y+n_x} f_{ss}^i (s', s_t) \right] & \left[ +2H_{2n_y+1,2n_y+n_x;2n_y+1,2n_y+n_x} f_{ss}^i (s', s_t) \right] & [D_{n_x+1,n_x+n_z} h_{ss} (s_t)] \\
+D_{n_x+1,n_x+n_z} g_{ss} (s_t)^T & \left[ H_{n_y+1,2n_y,n_y+1,2n_y+n_x} f_{ss}^i (s', s_t) \right] & \left[ +2H_{n_y+1,2n_y;2n_y+1,2n_y+n_x} f_{ss}^i (s', s_t) \right] & [D_{n_x+1,n_x+n_z} h_{ss} (s_t)] \\
+D_{n_x+1,n_x+n_z} h_{ss} (s_t)^T & \left[ H_{n_y+1,2n_y,n_y+1,2n_y+n_x} f_{ss}^i (s', s_t) \right] & \left[ +2H_{n_y+1,2n_y;2n_y+1,2n_y+n_x} f_{ss}^i (s', s_t) \right] & [D_{n_x+1,n_x+n_z} h_{ss} (s_t)] \\
\end{pmatrix}
\]
\[
\mathcal{H}_{n_{x}+1, n_{z}; n_{x}+n_{z}+1} G_{s_{ss}}^i (s_t) = \sum_{s'_{-1}}^n p_{s_{t}, s'} \times \left[ \begin{array}{c}
\mathcal{H}_{n_{y}; n_{y}+1} f_{s_{ss}}^i (s', s_t) \\
+ \mathcal{D}_{n_{y}; n_{y}+1} g_{s_{ss}} (s')
\end{array} \right] \begin{array}{c}
\mathcal{D}_{n_{x}+n_{z}+1} g_{s_{ss}} (s') \\
+ \mathcal{D}_{n_{x}+n_{z}+1} h_{s_{ss}} (s_t)
\end{array}
\right] + \mathcal{D}_{n_{x}+1, n_{x}+n_{z}} h_{s_{ss}} (s_t)^T + \mathcal{D}_{n_{x}+1, n_{x}+n_{z}} g_{s_{ss}} (s_t)^T
\]